

Sample Questions: Moment-generating functions

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1. Let X have a moment-generating function $M_X(t)$ and let a be a constant. Show $M_{aX}(t) = M_X(at)$.

$$\begin{aligned} M_{aX}(t) &= E\left(e^{atX}\right) = E\left(e^{X(at)}\right) \\ &= M_X(at) \end{aligned}$$

2. Let X have a moment-generating function $M_x(t)$ and let a be a constant. Show $M_{a+X}(t) = e^{at}M_x(t)$.

$$\begin{aligned}M_{a+X}(t) &= E\left(e^{(a+X)t}\right) \\&= E\left(e^{at+Xt}\right) = E\left(e^{at}e^{Xt}\right) \\&= e^{at}E\left(e^{Xt}\right) = e^{at}M_x(t)\end{aligned}$$

3. Let X and Y be independent, (continuous) random variables.
Show $M_{x+y}(t) = M_x(t) M_y(t)$.

$$\begin{aligned} M_{x+y}(t) &= E\left(e^{(x+y)t}\right) = E\left(e^{xt+yt}\right) \\ &= E\left(e^{xt} e^{yt}\right) \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{xt} e^{yt} f_{x,y}(x,y) dx dy \\ &\stackrel{\text{ind}}{=} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{xt} e^{yt} f_x(x) f_y(y) dx dy \\ &= \int_{-\infty}^{\infty} e^{yt} f_y(y) \left(\int_{-\infty}^{\infty} e^{xt} f_x(x) dx \right) dy \\ &= M_x(t) \int_{-\infty}^{\infty} e^{yt} f_y(y) dy = M_x(t) M_y(t) \end{aligned}$$

4. Let $X \sim N(0, 1)$. Calculate $M_x(t)$.

$$\begin{aligned}M_x(t) &= E(e^{xt}) = \int_{-\infty}^{\infty} e^{xt} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} dx \\&= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(x - 2xt + t^2 - t^2)} dx \\&= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(x^2 - 2xt + t^2)} e^{-\frac{1}{2}(-t^2)} dx \\&= e^{\frac{1}{2}t^2} \left[\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(x-t)^2} dx \right] \\&= e^{\frac{1}{2}t^2} \quad = 1\end{aligned}$$

5. Let $X \sim N(\mu, \sigma)$. Calculate $M_x(t)$. $Z = \frac{X - \mu}{\sigma}$ $M_z(t) = e^{\frac{1}{2}t^2}$

$$M_x(t) = E(e^{xt})$$

$$Z = \frac{X - \mu}{\sigma}$$

$$X = Z\sigma + \mu$$

$$= E(e^{(Z\sigma + \mu)t})$$

$$= E(e^{Z\sigma t} e^{\mu t}) = e^{\mu t} E(e^{Z(\sigma t)})$$

$$= e^{\mu t} M_z(\sigma t)$$

$$= e^{\mu t} e^{\frac{1}{2}\sigma^2 t^2}$$

$$= e^{\mu t + \frac{1}{2}\sigma^2 t^2}$$

$$M'(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2} \cdot (\mu + \frac{1}{2}\sigma^2 2t) \Big|_{t=0}$$

$$= \mu$$

6. Let $X \sim N(\mu, \sigma)$. Show $Y = \frac{X-\mu}{\sigma} \sim N(0, 1)$ using moment-generating functions.

$$\begin{aligned}
 M_Y(t) &= \cancel{E} E(e^{Yt}) = E(e^{(\frac{X-\mu}{\sigma})t}) \\
 &= E\left(e^{\left(\frac{Xt}{\sigma} - \frac{\mu t}{\sigma}\right)}\right) = E\left(e^{X\left(\frac{t}{\sigma}\right)} e^{-\frac{\mu t}{\sigma}}\right) \\
 &= e^{-\frac{\mu t}{\sigma}} E\left(e^{X\left(\frac{t}{\sigma}\right)}\right) = e^{-\frac{\mu t}{\sigma}} M_X\left(\frac{t}{\sigma}\right) \\
 &= e^{-\frac{\mu t}{\sigma}} e^{\mu\left(\frac{t}{\sigma}\right) + \frac{1}{2}\sigma^2\left(\frac{t}{\sigma}\right)^2} \\
 &= \cancel{e^{-\frac{\mu t}{\sigma}}} \cancel{e^{\frac{\mu t}{\sigma}}} e^{\frac{1}{2}\sigma^2\frac{t^2}{\sigma^2}} \\
 &= e^{\frac{1}{2}t^2} \quad \text{MGF of } N(0, 1)
 \end{aligned}$$

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<http://www.utstat.toronto.edu/~brunner/oldclass/256f18>

$$M_x(t) = e^{\mu t + \frac{1}{2} \sigma^2 t^2}$$

7. Let $X \sim N(\mu, \sigma)$. Find the distribution of $Y = a + bX$ using moment-generating functions.

$$\begin{aligned} M_Y(t) &= E(e^{tY}) = E(e^{(a+bX)t}) \\ &= E(e^{at} e^{X(bt)}) = e^{at} E(e^{X(bt)}) \\ &= e^{at} M_x(bt) = e^{at} e^{\mu(bt) + \frac{1}{2} \sigma^2 (bt)^2} \\ &= e^{(a+b\mu)t + \frac{1}{2} (b^2 \sigma^2) t^2} \end{aligned}$$

MGF of $N(a+b\mu, b\sigma)$ ✓

$$e^{\mu t + \frac{1}{2} \sigma^2 t^2}$$

8. Let $X_1 \sim N(\mu_1, \sigma_1)$ and $X_2 \sim N(\mu_2, \sigma_2)$ be independent. Find the distribution of $Y = aX_1 + bX_2 + c$.

$$M_Y(t) = E(e^{yt}) = E(e^{(aX_1 + bX_2 + c)t})$$

$$= E(e^{aX_1 t} e^{bX_2 t} e^{ct})$$

$$= e^{ct} E(e^{aX_1 t}) E(e^{bX_2 t})$$

independence of X_1 & X_2 &
hence of $aX_1 t$ & $bX_2 t$

$$= e^{ct} M_{aX_1}(t) M_{bX_2}(t)$$

$$= e^{ct} M_{X_1}(at) M_{X_2}(bt)$$

$$= e^{ct} e^{\mu_1 at + \frac{1}{2} \sigma_1^2 a^2 t^2} e^{\mu_2 bt + \frac{1}{2} \sigma_2^2 (bt)^2}$$

$$e^{(a\mu_1 + b\mu_2 + c)t + \frac{1}{2} (a^2 \sigma_1^2 + b^2 \sigma_2^2) t^2}$$

MGF of

$$N(a\mu_1 + b\mu_2 + c, \sqrt{a^2 \sigma_1^2 + b^2 \sigma_2^2})$$

9. Let $Z \sim N(0, 1)$ and let $Y = Z^2$. Find the distribution of Y . Recall that the MGF of a chi-squared random variable is $M(t) = (1 - 2t)^{-\frac{\nu}{2}}$.

$$\begin{aligned}
 M_Y(t) &= E(e^{Yt}) = E(e^{Z^2 t}) \\
 &= \int_{-\infty}^{\infty} e^{z^2 t} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2} dz \\
 &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} z^2 (1-2t)} dz \\
 &= (1-2t)^{-\frac{1}{2}} \left(\frac{1}{(1-2t)^{\frac{1}{2}} \sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{z^2}{(1-2t)^{-1}}} dz \right) \\
 &= (1-2t)^{-\frac{1}{2}} \text{ like } (1-2t)^{-\frac{\nu}{2}}
 \end{aligned}$$

Integral of normal = 1

$\frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}$

MGF of $\chi^2(1)$

10. Independently for $i = 1, \dots, n$, let $Y_i \sim \chi^2(\nu_i)$. Find the distribution of $W = \sum_{i=1}^n Y_i$.

$$\begin{aligned} M_W(t) &= M_{\sum Y_i}(t) \stackrel{\text{ind}}{=} \prod_{i=1}^n M_{Y_i}(t) \\ &= \prod_{i=1}^n (1-2t)^{-\frac{\nu_i}{2}} = (1-2t)^{-\frac{\sum_{i=1}^n \nu_i}{2}} \end{aligned}$$

$$\text{MGF of } \chi^2\left(\sum_{i=1}^n \nu_i\right)$$

11. Independently for $i = 1, \dots, n$, let $X_i \sim N(\mu_i, \sigma_i)$. What is the distribution of $Y = \sum_{i=1}^n \left(\frac{X_i - \mu_i}{\sigma_i} \right)^2$? Justify your answer.

$\chi^2(n)$ because

$$\frac{X_i - \mu_i}{\sigma_i} \sim N(0, 1)$$

Square of standard normal is $\chi^2(1)$

Sum of independent χ^2 's is

$$\chi^2(\sum \nu_i)$$

$$\text{so } \chi^2(n)$$

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