



Sample Questions: Expected Value, Variance and Covariance

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1. Let X have a continuous uniform distribution on (a, b) . Calculate $E(X)$.

$$\begin{aligned}
 E(X) &= \int_a^b x \frac{1}{b-a} dx = \frac{1}{b-a} \left. \frac{x^2}{2} \right|_a^b \\
 &= \frac{1}{2} \frac{1}{b-a} (b^2 - a^2) = \frac{1}{2} \frac{1}{\cancel{b-a}} (\cancel{b-a})(b+a) \\
 &= \frac{b+a}{2}
 \end{aligned}$$

2. Let $X \sim \text{Poisson}(\lambda)$. Calculate $E(X)$.

$$E(X) = \sum_{x=0}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = \sum_{x=1}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x(x-1)!}$$

$$\begin{array}{l} k = x - 1 \\ x = k + 1 \end{array} \quad \begin{array}{c|c} x & k = x - 1 \\ \hline \infty & \infty \\ 1 & 0 \end{array}$$

$$= \sum_{k=0}^{\infty} \frac{e^{-\lambda} \lambda^{k+1}}{k!} = \lambda \sum_{k=0}^{\infty} \frac{e^{-\lambda} \lambda^k}{k!}$$

$$= \lambda$$

$= 1$

3. Let the continuous random variable X have density $f(x) = \begin{cases} \frac{1}{x^2} & \text{for } x \geq 1 \\ 0 & \text{otherwise} \end{cases}$

(a) Verify that $f(x)$ integrates to one.

$$\begin{aligned} \int_1^{\infty} x^{-2} dx &= \frac{x^{-1}}{-1} \Big|_1^{\infty} = (-1) \frac{1}{x} \Big|_1^{\infty} \\ &= (-1) (0 - 1) = 1 \end{aligned}$$

(b) Calculate $E(X)$.

$$\begin{aligned} E(X) &= \int_1^{\infty} x \frac{1}{x^2} dx = \int_1^{\infty} \frac{1}{x} dx \\ &= \ln x \Big|_1^{\infty} = \infty \end{aligned}$$

$E(X)$ does not exist

4. Let $X \sim N(\mu, \sigma)$. Calculate $E(X)$.

$$E(X) = \int_{-\infty}^{\infty} x \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2} dx$$

$$z = \frac{x-\mu}{\sigma} \quad dz = \frac{1}{\sigma} dx \quad \begin{array}{c|c} x & z \\ \hline \infty & \infty \\ -\infty & -\infty \end{array}$$

$$x - \mu = z\sigma \\ x = z\sigma + \mu$$

$$= \int_{-\infty}^{\infty} (z\sigma + \mu) \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2} dz$$

$$= \sigma \int_{-\infty}^{\infty} z \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2} dz + \mu \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2} dz}_{=1}$$

$$= \mu + \sigma \int_{-\infty}^0 z \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2} dz + \sigma \int_0^{\infty} z \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2} dz$$

$$u = \frac{1}{2} z^2 \quad du = z dz$$

x	u
0	0
$-\infty$	∞

x	$u = \frac{1}{2} z^2$
∞	∞
0	0

$$= \mu + \frac{\sigma}{\sqrt{2\pi}} \underbrace{\int_{\infty}^0 e^{-u} du}_{=-1} + \frac{\sigma}{\sqrt{2\pi}} \underbrace{\int_0^{\infty} e^{-u} du}_{=1} = \mu$$

5. Let X have a binomial distribution with parameters n and p . Calculate $E(X)$.

$$E(X) = \sum_{x=0}^n x \binom{n}{x} p^x (1-p)^{n-x}$$

$$= \sum_{x=1}^n x \frac{n!}{x! (n-x)!} p^x (1-p)^{n-x}$$

$$k = x-1$$

$$x = k+1$$

$$\begin{array}{c|c} x & k = x-1 \\ \hline n & n-1 \\ \hline 1 & 0 \end{array}$$

$$= \sum_{k=0}^{n-1} \frac{n(n-1)!}{k! (n-(k+1))!} p^{k+1} (1-p)^{n-(k+1)}$$

$$= np \sum_{k=0}^{n-1} \frac{(n-1)!}{k! (n-1-k)!} p^k (1-p)^{n-1-k}$$

$$= (p + (1-p))^n$$

$$= np$$

$$= 1$$

Toss fair coin 100 times. How many heads would you expect?

$$np = 100 \cdot \frac{1}{2} = 50$$

6. Let X have a Gamma distribution with parameters α and λ . Calculate $E(X^k)$.

$$E(X^k) = \int_0^{\infty} x^k \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda x} x^{\alpha-1} dx$$

$$= \frac{\lambda^\alpha}{\Gamma(\alpha)} \frac{\Gamma(\alpha+k)}{\lambda^\alpha \lambda^k} \int_0^{\infty} \frac{\lambda^{\alpha+k}}{\Gamma(\alpha+k)} e^{-\lambda x} x^{(\alpha+k)-1} dx$$

$$= \frac{\Gamma(\alpha+k)}{\lambda^k \Gamma(\alpha)}$$

T

= 1

$$\Gamma(\alpha+1) = \alpha \Gamma(\alpha)$$

For or $k=1$

$$\frac{\Gamma(\alpha+1)}{\lambda \Gamma(\alpha)}$$

$$= \frac{\alpha \cancel{\Gamma(\alpha)}}{\lambda \cancel{\Gamma(\alpha)}}$$

$$= \left(\frac{\alpha}{\lambda} \right)$$

7. Let X and Y be independent (continuous) random variables.
Show $E(XY) = E(X)E(Y)$.

$$E(XY) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{XY}(x, y) dx dy$$

ind
 \Downarrow
 \Downarrow

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_X(x) f_Y(y) dx dy$$

$$\int_{-\infty}^{\infty} y f_Y(y) \left(\int_{-\infty}^{\infty} x f_X(x) dx \right) dy$$

$E(X)$

$$= E(X) \int_{-\infty}^{\infty} y f_Y(y) dy = E(X)E(Y)$$

$$E(a+X) = a + E(X) = a + \overset{E(X)}{\mu}$$

8. Prove $\text{Var}(a+X) = \text{Var}(X)$.

$$\text{Var}(a+X) = E\left\{(a+X - (a+\mu))^2\right\}$$

$$= E\left\{(\cancel{a}+X - \cancel{a} - \mu)^2\right\}$$

$$= E\left\{(X - \mu)^2\right\} = \text{Var}(X)$$

$$E(bX) = b\mu$$

9. Prove $\text{Var}(bX) = b^2\text{Var}(X)$.

$$\text{Var}(bX) = E \left\{ (bX - b\mu)^2 \right\}$$

$$= E \left\{ (b(X - \mu))^2 \right\} = E \left\{ b^2 (X - \mu)^2 \right\}$$

$$= b^2 E \left\{ (X - \mu)^2 \right\} = b^2 \text{Var}(X) \quad \checkmark$$

10. Show $\text{Var}(X) = E(X^2) - [E(X)]^2$.

$$E(X) = \mu$$

$$\text{Var}(X) = E\{(X - \mu)^2\}$$

$$= E\{X^2 - 2X\mu + \mu^2\}$$

$$= E(X^2) - 2\mu E(X) + E(\mu^2)$$

$$= E(X^2) - 2\mu\mu + \mu^2$$

$$= E(X^2) - \mu^2 = E(X^2) - [E(X)]^2$$



11. Let $X \sim \text{Uniform}(0,1)$. Calculate $\text{Var}(X)$.

$$\text{Var}(X) = E(X^2) - (E(X))^2$$

$$E(X) = \int_0^1 x \cdot 1 dx = \frac{x^2}{2} \Big|_0^1 = \frac{1}{2}$$

$$E(X^2) = \int_0^1 x^2 dx = \frac{x^3}{3} \Big|_0^1 = \frac{1}{3}$$

$$E(X^2) - (E(X))^2 = \frac{1}{3} - \left(\frac{1}{2}\right)^2 = \frac{1}{3} - \frac{1}{4}$$

$$= \frac{4}{12} - \frac{3}{12} = \frac{1}{12}$$

12. Let X have density e^{-x} for $x \geq 0$ and zero otherwise. Calculate $\text{Var}(X)$.

$$\Gamma(\alpha) = \int_0^{\infty} e^{-t} t^{\alpha-1} dt, \quad \Gamma(\alpha+1) = \alpha \Gamma(\alpha)$$
$$\Gamma(1) = \int_0^{\infty} e^{-t} dt = 1$$

$$E(X) = \int_0^{\infty} x e^{-x} dx = \int_0^{\infty} e^{-x} x^{2-1} dx = \Gamma(2)$$
$$= 1 \Gamma(1) = 1$$

$$E(X^2) = \int_0^{\infty} x^2 e^{-x} dx = \int_0^{\infty} e^{-x} x^{3-1} dx$$
$$= \Gamma(3) = 2 \Gamma(2) = 2$$

$$\text{Var}(X) = E(X^2) - (E(X))^2 = 2 - 1^2 = 1$$

13. Let $X \sim N(\mu, \sigma)$. Calculate $\text{Var}(X)$.

$$\Gamma(\alpha) = \int_0^{\infty} e^{-t} t^{\alpha-1} dt$$

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

$$\Gamma(\alpha+1) = \alpha \Gamma(\alpha)$$

$$\text{Var}(X) = E\{(X-\mu)^2\}$$

$$= \sigma^2 \int_{-\infty}^{\infty} \frac{(x-\mu)^2}{\sigma^2} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx$$

$$z = \frac{x-\mu}{\sigma}$$

$$dz = \frac{1}{\sigma} dx$$

$$= \sigma^2 \int_{-\infty}^{\infty} z^2 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}z^2} dz$$

symmetric around $z=0$

$$\begin{array}{c|c} x & z \\ \hline \infty & \infty \\ \hline -\infty & -\infty \end{array}$$

$$= \frac{2\sigma^2}{\sqrt{2\pi}} \int_0^{\infty} z e^{-\frac{1}{2}z^2} z dz$$

$$\begin{array}{c} u = \frac{1}{2}z^2 \\ \frac{z}{2} = \frac{u}{z} \\ \frac{z}{2} = \frac{u}{z} \\ \hline \infty \quad \infty \\ \hline 0 \quad 0 \end{array} \quad du = z dz$$

$$= \frac{2\sigma^2}{\sqrt{2}\sqrt{\pi}} \int_0^{\infty} \sqrt{2} u^{\frac{1}{2}} e^{-u} du$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \int_0^{\infty} e^{-u} u^{\frac{3}{2}-1} du = \frac{2\sigma^2}{\sqrt{\pi}} \Gamma\left(\frac{3}{2}\right)$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \cdot \frac{1}{2} \Gamma\left(\frac{1}{2}\right) = \sigma^2$$

14. The discrete random variables x and y have joint distribution

	$x = 1$	$x = 2$	$x = 3$	
$y = 1$	3/12	1/12	3/12	7/12
$y = 2$	1/12	3/12	1/12	5/12

(a) What is $E(X|Y = 1)$? $\frac{4}{12}$ $\frac{4}{12}$ $\frac{4}{12}$

$$= \sum_x x P(X=x|Y=1)$$

$$= 1 \cdot \frac{3/12}{7/12} + 2 \cdot \frac{1/12}{7/12} + 3 \cdot \frac{3/12}{7/12}$$

$$= \frac{1}{7} (3 + 2 + 9) = \frac{14}{7} = 2$$

(b) What is $E(Y^2|X = 2)$?

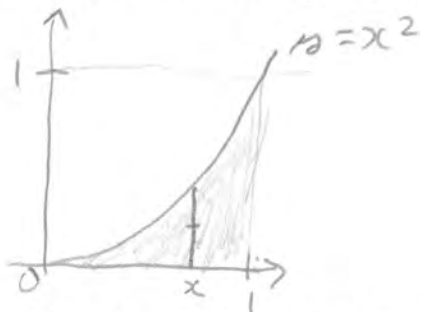
$$= \sum_{y_2} y_2^2 P(Y=y_2|X=2)$$

$$= 1^2 \frac{1/12}{4/12} + 2^2 \frac{3/12}{4/12}$$

$$= \frac{1}{4} + \frac{12}{4} = \frac{13}{4}$$

	$x = 1$	$x = 2$	$x = 3$
$y = 1$	3/12	1/12	3/12
$y = 2$	1/12	3/12	1/12

15. Let $f_{x,y}(x,y) = 3$ for $0 < x < 1$ and $0 < y < x^2$, and zero otherwise.



- (a) Using $f_x(x) = 3x^2$ for $0 < x < 1$, what is $f_{y|x}(y|x)$? Don't forget the support.

$$f_{y|x}(y|x) = \begin{cases} \frac{3}{3x^2} & 0 \leq y \leq x^2 \\ 0 & \text{otherwise} \end{cases}$$

where $0 < x < 1$

- (b) Find $E(Y|X = x)$, where x is a fixed constant between zero and one.

$$\begin{aligned} E(Y|X=x) &= \int_0^{x^2} y \frac{1}{x^2} dy = \frac{1}{x^2} \int_0^{x^2} y dy \\ &= \frac{1}{x^2} \left. \frac{y^2}{2} \right|_0^{x^2} = \frac{1}{x^2} \frac{1}{2} x^4 = \left(\frac{x^2}{2} \right) \end{aligned}$$

(c) Find $E(Y)$ by double expectation. $E(Y) = E(E(Y|X))$

$$= E\left(\frac{X^2}{2}\right) = \int_0^1 \frac{x^2}{2} \cdot 3x^2 dx$$

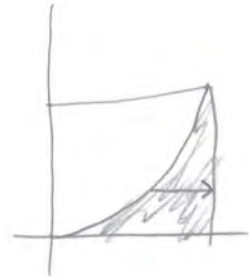
$$= \frac{3}{2} \int_0^1 x^4 dx = \frac{3}{2} \frac{x^5}{5} \Big|_0^1 = \frac{3}{10}$$

(d) Using $f_y(y) = 3(1 - y^{1/2})$ for $0 < y < 1$, calculate $E(Y)$ directly.

$$E(Y) = \int_0^1 3y(1 - y^{1/2}) dy$$

$$= \int_0^1 3(y - y^{3/2}) dy$$

$$= 3y^2 \Big|_0^1 - \frac{3y^{5/2}}{5/2} \Big|_0^1 = 3\left(\frac{1}{2} - \frac{2}{5}\right) = 3\left(\frac{5}{10} - \frac{4}{10}\right) = \frac{3}{10}$$



16. Show that $Cov(X, Y) = E(XY) - E(X)E(Y)$

$$\begin{aligned} Cov(X, Y) &\stackrel{\text{def}}{=} E\{(X - \mu_x)(Y - \mu_y)\} \\ &= E\{XY - X\mu_y - \mu_x Y + \mu_x \mu_y\} \\ &= E\{XY\} - E(X)\mu_y - \mu_x E(Y) + \mu_x \mu_y \\ &= E(XY) - \mu_x \mu_y - \mu_x \mu_y + \mu_x \mu_y \\ &= E(XY) - E(X)E(Y) \end{aligned}$$

17. Show that if X and Y are independent, $Cov(X, Y) = 0$.

Have seen if X & Y are independent $E(XY) = E(X)E(Y)$

$$\begin{aligned} \text{So } Cov(X, Y) &= E(XY) - E(X)E(Y) \\ &\stackrel{\text{ind}}{=} E(X)E(Y) - E(X)E(Y) = 0 \end{aligned}$$

18. The discrete random variables x and y have joint distribution

	$x = 1$	$x = 2$	$x = 3$	
$y = 1$	$3/12$	$1/12$	$3/12$	$7/12$
$y = 2$	$1/12$	$3/12$	$1/12$	$5/12$

(a) Find $Cov(X, Y)$.

$$E(X) = 2, \quad E(Y) = \sum_y y P(y) = 1 \cdot \frac{7}{12} + 2 \cdot \frac{5}{12} = \frac{17}{12}$$

	1	2	3
1	1	2	3
2	2	4	6

$$E(XY) = \frac{1}{12} (3 + 2 + 9 + 2 + 6) = \frac{34}{12}$$

$$Cov(X, Y) = E(XY) - E(X)E(Y) = \frac{34}{12} - \frac{34}{12} = 0$$

(b) Are X and Y independent?

$$P(1, 1) \neq P_X(1) P_Y(1) = \frac{1}{3} \cdot \frac{7}{12} = \frac{7}{36}$$

$\frac{3}{12}$

No, NOT ind.

19. Show $\text{Var}(aX + bY) = a^2\text{Var}(X) + b^2\text{Var}(Y) + 2ab\text{Cov}(X, Y)$

$$E(aX + bY) = a\mu_x + b\mu_y$$

$$\text{Var}(aX + bY) \stackrel{\text{def}}{=} E\left\{\left(aX + bY - (a\mu_x + b\mu_y)\right)^2\right\}$$

$$= E\left\{\left(a(X - \mu_x) + b(Y - \mu_y)\right)^2\right\}$$

$$= E\left\{a^2(X - \mu_x)^2 + b^2(Y - \mu_y)^2 + 2ab(X - \mu_x)(Y - \mu_y)\right\}$$

$$= a^2E(X - \mu_x)^2 + b^2E(Y - \mu_y)^2 + 2abE\left\{(X - \mu_x)(Y - \mu_y)\right\}$$

$$= a^2\text{Var}(X) + b^2\text{Var}(Y) + 2ab\text{Cov}(X, Y)$$

20. Prove that $Var(\sum_{i=1}^n a_i X_i) = \sum_{i=1}^n a_i^2 Var(X_i) + 2 \sum_{\substack{i=1 \\ i \neq j}}^{n-1} \sum_{j=i+1}^n a_i a_j Cov(X_i, X_j)$.

$$E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i \mu_i$$

$$Var\left(\sum_{i=1}^n a_i X_i\right) \stackrel{def}{=} E\left\{\left(\sum_{i=1}^n a_i X_i - \sum_{i=1}^n a_i \mu_i\right)^2\right\}$$

$$= E\left\{\left(\sum_{i=1}^n a_i (X_i - \mu_i)\right)^2\right\}$$

	1	2	...	n	
1		x	y	...	x
2			x	...	y
...					
n					x

$$= E\left\{\left(a_1(X_1 - \mu_1) + a_2(X_2 - \mu_2) + \dots + a_n(X_n - \mu_n)\right)^2\right\}$$

$$= E\left\{\sum_{i=1}^n \sum_{j=1}^n a_i a_j (X_i - \mu_i)(X_j - \mu_j)\right\}$$

$$= \sum_{i=1}^n \sum_{j=1}^n a_i a_j E\left\{(X_i - \mu_i)(X_j - \mu_j)\right\}$$

$$= \sum_{i=1}^n a_i^2 Var(X_i) + \sum_{\substack{i=1 \\ i \neq j}}^n \sum_{j=i+1}^n a_i a_j Cov(X_i, X_j) \quad \checkmark$$

This handout was prepared by Jerry Brunner, Department of Mathematical and Computational Sciences, University of Toronto. It is licensed under a Creative Commons Attribution - ShareAlike 3.0 Unported License. Use any part of it as you like and share the result freely. The L^AT_EX source code is available from the course website:

<http://www.utstat.toronto.edu/~brunner/oldclass/256f18>