

May 2013

Ph.D. COMPREHENSIVE EXAMINATIONS  
DEPARTMENT OF STATISTICS  
UNIVERSITY OF TORONTO

THEORETICAL STATISTICS COMPREHENSIVE EXAMINATION

1. *Attempt all questions* (total # of questions = 7).  
It is not necessary to completely solve every problem to achieve a good performance.  
*Emphasize what you do know.*
2. Please work neatly and legibly.
3. *Start each question in a new book, with your name and the number of the question on the front cover.* If there is more than one book for a question, then also indicate which is the first book and which second, e.g., Jane Smith, Question 5, Book 1 of 2.
4. The questions are not in any special order, nor are they all of equal difficulty.
5. The problems may be improperly phrased or may contain a misprint. Should this happen, reflect it in your discussion. Faculty members are *not* available to answer questions during the exam.
6. You are *not* permitted any aids (e.g., books, notes, etc.) aside from a single non-programmable calculator.
7. Probability tables that may be useful are appended at the end of the exam paper, after all questions.
8. Good luck!

1. Suppose that  $x = (x_1, \dots, x_n) \stackrel{i.i.d.}{\sim} N(\mu, \sigma^2)$  and  $\mu | \sigma^2 \sim N(\mu_0, \tau_0^2 \sigma^2)$ ,  $1/\sigma^2 \sim \text{Gamma}(\alpha_0, \beta_0)$  with  $\mu_0, \tau_0^2, \alpha_0$  and  $\beta_0$  fixed.
  - (a) Show that  $T(x) = (\bar{x}, \|x - \bar{x}1\|^2)$  is a minimal sufficient statistic and determine its distribution.
  - (b) Determine the posterior distribution of  $(\mu, \sigma^2)$ .
  - (c) Determine the posterior distribution of  $\mu$  and the form of hpd intervals for  $\mu$ .
  - (d) Determine the posterior predictive distribution of a future  $y \sim N(\mu, \sigma^2)$ .
  - (e) Discuss how  $\mu_0, \tau_0^2, \alpha_0$  and  $\beta_0$  are to be chosen in an application.
  
2. Let  $X_n \sim \text{Multinomial}(n, p_n)$  and  $Y_n \sim \text{Multinomial}(c_n n, q_n)$ , where  $X_n$  and  $Y_n$  are independent with the same dimension  $k$ ,  $c_n$  is chosen so that  $c_n n$  is an integer with  $c_n \rightarrow c > 0$  as  $n \rightarrow \infty$ . Assume that  $\sqrt{n}(p_n - p) \rightarrow 0$  and  $\sqrt{n}(q_n - q) \rightarrow 0$ .
  - (a) Derive the asymptotic distribution of a properly scaled form of  $\left(\frac{X_n}{n} - \frac{Y_n}{c_n n}\right)$ .
  - (b) Construct an appropriate statistic  $W_n$  in the quadratic form of the quantity in (a), such that  $W_n$  asymptotically follows a Chi-square distribution under the null hypothesis that  $p = q$ . Justify all steps and specify the limiting distribution. You may use the fact on projection matrix.
  
3. Suppose that  $y \sim N_n(X\beta, \sigma^2 I)$  where  $X \in R^{n \times k}$  is of rank  $k$  and known, while  $(\beta, \sigma^2) \in R^k \times (0, \infty)$ .
  - (a) Determine the MLE  $(\hat{\beta}, \hat{\sigma}^2)$  of  $(\beta, \sigma^2)$  and its joint distribution.
  - (b) Show that  $\hat{\beta}$  is an optimal unbiased estimator of  $\beta$  and explain what the criterion for optimality is.
  - (c) Determine the optimal unbiased estimator of  $\sigma^2$ .
  
4. Let  $X_i \stackrel{i.i.d.}{\sim} \text{Bernoulli}(1/2)$  and  $\{Y_i | X_i = c\} \stackrel{i.i.d.}{\sim} \text{Poisson}(c\theta)$ , and one only can observe  $Y_i$ 's.
  - (a) Find an estimator of  $\theta$  based on  $\bar{Y}_n = \sum_{i=1}^n Y_i$  and its asymptotic distribution.
  - (b) Find an asymptotically efficient estimator of  $\theta$ , and specify the asymptotic distribution of this estimator.
  
5. Let  $g : R^r \rightarrow R$  be symmetric in its arguments,  $F \in \mathcal{F} = \{G : \int \cdots \int |g^2(y_1, \dots, y_r)| \prod_{i=1}^r dG(y_i) < \infty\}$ , consider the V-functional and the corresponding estimator,

$$T(f) = \int \cdots \int g(y_1, \dots, y_r) \prod_{i=1}^r dF(y_i), \quad T(F_n) = n^{-r} \sum_{i_1=1}^n \cdots \sum_{i_r=1}^n g(X_{i_1}, \dots, X_{i_r}),$$

where  $\{X_1, \dots, X_n\}$  is a random sample from  $F$ .

- (a) Calculate the influence function  $T'(x; F)$ .
- (b) Drive the asymptotic distribution of the properly scaled  $T(F_n)$ . Note that it is necessary to show the remaining term is asymptotically negligible.
6. Suppose that  $X_1, \dots, X_n$  are independently and identically distributed with the probability density function  $f_{\theta, j}$ , where  $f_{\theta, 1}$  is the density of  $N(0, \theta^2)$  and  $f_{\theta, 2} = (2\theta)^{-1} e^{-|x|/\theta}$ ,  $\theta > 0$ ,  $j = 1, 2$ .
- (a) Obtain the MLE of  $(\theta, j)$ .
- (b) Is the MLE of  $j$  obtained in (a) is consistent? Justify.
- (c) Show that the MLE of  $\theta$  is consistent and find its asymptotic distribution.
7. Suppose  $G$  is a symmetry group of a model  $\{f_\theta : \theta \in \Theta\}$  with sample space  $\mathcal{X}$ .
- (a) If  $S : \mathcal{X} \rightarrow \mathcal{S}$  is onto and equivariant under  $G$ , then show that this implies an action of  $G$  on  $\mathcal{S}$  given by  $T_g^S s = S(T_g x)$  when  $S(x) = s$ .
- (b) If  $\Psi : \Theta \rightarrow \Psi$  is onto (so  $\Psi$  denotes both the map and its range), define what it means for  $\Psi(\theta)$  to be equivariant under  $G$  and give the action  $T_g^*$  this induces on  $\Psi$ , namely, what does  $T_g^* \psi$  equal.
- (c) Give the full definition of what it means for a decision problem to be invariant under  $G$ .
- (d) Suppose  $\{f_{(\mu, \sigma)} : \mu \in R^1, \sigma > 0\}$  is a location-scale family on  $R^n$ ,  $\Psi(\mu, \sigma) = \mu$  and  $L((\mu, \sigma), \psi) = (\mu - \psi)^2 / \sigma^2$ . Show that this decision problem is invariant under the location-scale group.
- (e) For  $g \in G$  and  $\delta \in \mathcal{D}$ , where  $\mathcal{D}$  is the set of all decision functions for a problem, show that  $T_g^{**} \delta$  defined by  $T_g^{**} \delta(x, A) = \delta(T_g^{-1} x, T_g^{*-1} A)$  defines an action on  $\mathcal{D}$ .
- (f) Prove that a nonrandomized  $\delta$  is invariant if and only if the corresponding estimator  $d$  is equivariant under  $G$ .
- (g) Prove that if  $G$  leaves the decision problem invariant, then

$$R(\bar{T}_g \theta, T_g^{**} \delta) = R(\theta, \delta)$$

for every  $g, \theta$  and  $\delta$ . Conclude that, if  $\delta$  is invariant then  $\delta$  has constant risk on orbits of  $G$  acting on  $\Theta$ . What do you conclude if  $G$  is transitive on  $\Theta$ ? Explain why this transitivity simplifies the problem of obtaining the optimal invariant decision function.

- (h) Suppose that  $G$  leaves the decision problem invariant, acts transitively on  $\Theta$  and acts freely on  $\mathcal{X}$ . Let  $[\cdot]$  be a transformation variable and  $D$  be the corresponding maximal invariant. Prove that the optimal invariant estimator  $d$  is obtained by minimizing

$$\int_G L(\theta_0, T_g^* d(D(x))) P_{\theta_0[\cdot]}(dg | D)(D(x))$$

for  $d(D(x))$  where  $\theta_0$  is any fixed value of the full parameter. This estimator is called the *Pitman estimator*.

- (i) Obtain the Pitman estimator of  $\mu$  for the problem described in (d) when  $f_{(0,1)}$  is the  $N_n(0, I)$  density.