

STA3000 Sufficiency and Unbiased Estimation

This is treated in TPE at length, and also in CH §8.4. If $T = t(Y)$ is unbiased for θ , and $S = s(Y)$ is sufficient for θ (in the model $f(y; \theta)$, etc.), then $W = E(T | S)$ is unbiased for θ and has smaller variance than T .

To see this,

$$\begin{aligned}\theta = E(T; \theta) &= E\{E(T | S); \theta\} = E(W; \theta) \\ \text{var}(T; \theta) &= \text{var}\{E(T | S); \theta\} + E\{\text{var}(T | S); \theta\} = \text{var}(W; \theta) + E\{\text{var}(T | S); \theta\},\end{aligned}$$

where by the sufficiency of S , neither $E(T | S)$ nor $\text{var}(T | S)$ depend on θ . So W is unbiased, with smaller variance unless $\text{var}(T | S) = 0$.

This minimum variance unbiased estimator is also unique, under the assumption that S is a complete sufficient statistic. A sufficient statistic is complete if $E\{h(S); \theta\} = 0$ for every value θ implies $h(S) = 0$. (A sufficient statistic is boundedly complete if the same holds for bounded functions $h(\cdot)$). This is mainly/only used in establishing uniqueness of optimal tests and estimators based on sufficient statistics; it ensures that if we have two unbiased estimators of θ we end up at the same conditional estimator.

The Lehmann-Scheffé theorem says that if a sufficient statistic is boundedly complete it is minimal sufficient.

In i.i.d. sampling from an exponential family, the minimal sufficient statistic $T = t(Y)$ is complete if and only if the dimension of T is the same as the dimension of θ .