

Directional tests for vector parameters

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Parametric models and likelihood

- ▶ model $f(y; \theta)$, $\theta \in \mathbb{R}^p$
- ▶ data $y = (y_1, \dots, y_n)$ independent observations
- ▶ log-likelihood function $\ell(\theta; y) = \log f(y; \theta)$
- ▶ parameter of interest $\theta = (\psi, \lambda), \quad \psi \in \mathbb{R}^d$

- ▶ likelihood inference $w(\psi) = 2\{\ell(\hat{\psi}, \hat{\lambda}) - \ell(\psi, \hat{\lambda}_\psi)\}$

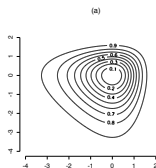
Example: 2×3 contingency table

- ▶ contingency table on activity amongst psychiatric patients (Everitt, 1992 CH)

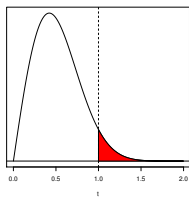
	Affective disorders	Schizophrenics	Neurotics
Retarded	12	13	5
Not retarded	18	17	25

- ▶ model: log-linear $y \sim \text{Poisson}$, $\log\{E(y)\} = X\theta$, $\theta \in \mathbb{R}^6$
- ▶ log-likelihood $\ell(\theta; y) = \theta'X'y - 1'e^{X\theta} = \theta's - c(\theta)$
- ▶ $\theta = (\psi, \lambda)$ $\psi \in \mathbb{R}^2, \lambda \in \mathbb{R}^4$
- ▶ (ψ_1, ψ_2) interaction parameters
- ▶ $H_0 : \psi = \psi_0 = (0, 0)$ independence
- ▶ log-likelihood $\ell(\psi, \lambda; y) = \psi' \underline{s}_1 + \lambda' \underline{s}_2 - c(\psi, \lambda)$

Testing $\psi = 0$



$$w(\psi_0) \sim \chi_2^2 \quad p\text{-value } 0.047$$



directional p -value **0.050**

'exact' conditional p -value **0.051**

Directional tests

- ▶ log-likelihood

$$\ell(\boldsymbol{\psi}, \boldsymbol{\lambda}) = \boldsymbol{\psi}' \underline{\mathbf{s}}_1 + \boldsymbol{\lambda}' \underline{\mathbf{s}}_2 - c(\boldsymbol{\psi}, \boldsymbol{\lambda})$$

- ▶ Step 1: get rid of nuisance parameter
use conditional density of $\underline{\mathbf{s}}_1$, given $\underline{\mathbf{s}}_2$
- ▶ Step 2: measure the **directed departure** from H_0
 \mathbf{s}_ψ : expected value of $\underline{\mathbf{s}}_1$, under H_0

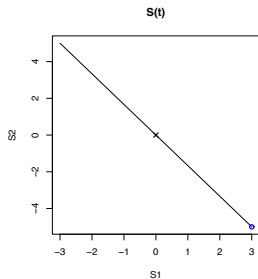
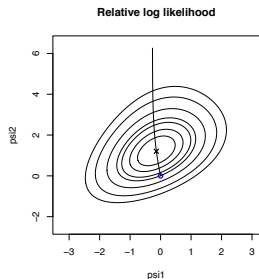
\mathbf{s}^0 : observed value of $\underline{\mathbf{s}}_1$

line through these two points

- ▶ $\mathcal{L}^* = t\mathbf{s}^0 + (1 - t)(\mathbf{s}_\psi - \mathbf{s}^0), \quad t \in \mathbb{R}$

... directional tests

$$\mathcal{L}^* = ts^0 + (1 - t)s_\psi$$



○ null hypothesis of independence $t = 0$
X observed value of s $t = 1$

$$p\text{-value} = \frac{\int_1^\infty g(t)dt}{\int_0^\infty g(t)dt}$$

like a 2-sided p -value

Pr (response $>$ observed | response $>$ 0)

... directional p -value

- ▶ p -value = $\frac{\int_1^\infty g(t) dt}{\int_0^\infty g(t) dt} = \frac{\int_1^\infty t^{d-1} f\{s(t); \psi\} dt}{\int_0^\infty t^{d-1} f\{s(t); \psi\} dt}$
- ▶ sample space, dimension p ($= 6$), $f(s; \theta)$
- ▶ reduced space \mathcal{L}_ψ , dimension d ($= 2$), $f(\underline{s}_1; \psi)$
- ▶ line \mathcal{L}^* , dimension 1, $f\{s_1(t); \psi\}$
- ▶ directed departure – magnitude, given direction
polar coordinates $\rightarrow t^{d-1}$

Simplifications

- ▶ 1: $\mathcal{L}^* \subset \mathcal{L}_\psi \subset \mathbb{R}^p$
- ▶ 2: ratio of two integrals – drop any terms that don't depend on t
- ▶ 3: saddle point approximation to $f(s; \theta)$ gives a very accurate starting point

Aside: From full likelihood to density along the line

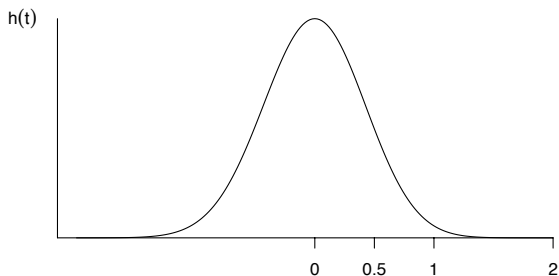
- ▶ suppose log-likelihood function $l(\varphi; \mathbf{s}) = \varphi' \mathbf{s} + l(\varphi; \mathbf{s}^0)$
- ▶ such as in exponential families
- ▶ $\varphi = \varphi(\theta); \quad \theta = (\psi, \lambda) = (\psi(\varphi), \lambda(\varphi))$

- ▶ double saddle point approximation gives

$$\hat{f}(\mathbf{s}; \psi) \doteq c \exp \left[l\{\varphi(\hat{\theta}_{\psi}^0); \mathbf{s}\} - l\{\hat{\varphi}(\mathbf{s}); \mathbf{s}\} \right] \frac{|j_{\varphi\varphi}\{\hat{\varphi}(\mathbf{s}); \mathbf{s}\}|^{-1/2}}{|j_{(\lambda\lambda)}\{\varphi(\hat{\theta}_{\psi}^0); \mathbf{s}\}|^{-1/2}}$$

- ▶ $\hat{\theta}_{\psi}^0 = (\psi, \hat{\lambda}_{\psi}^0)$
- ▶ $\hat{\varphi}(\mathbf{s})$ from $l_{\varphi}(\varphi; \mathbf{s}) = 0$
- ▶ uses only likelihood for the full model, plus constrained maximum likelihood estimator
- ▶ density along the line: use $f\{\mathbf{s}(t); \psi\}, t \in \mathbb{R}$
- ▶ $\hat{f}(\mathbf{s}; \psi)$ approximates the marginal likelihood (Kalbfleisch & Sprott, 1970, 1974)

2×3 table


$$t = 0$$

10	10	10
20	20	20

$$t = 0.5$$

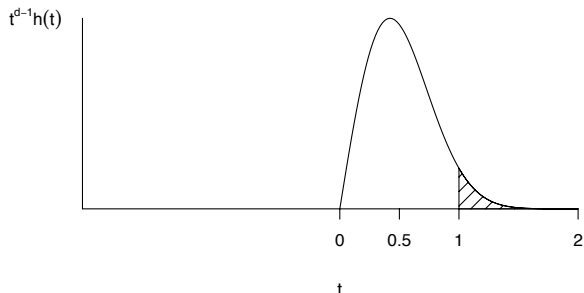
11.0	11.5	7.5
19.0	18.5	22.5

$$t = 1$$

12	13	5
18	17	25

$$t = 2$$

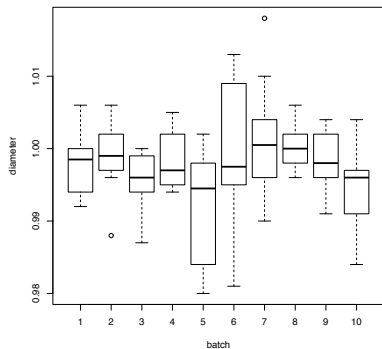
14	16	0
16	14	30



Simulations, 2×3 table

Nominal	0.010	0.025	0.050	0.100	0.250	0.500
LRT	0.011	0.028	0.055	0.107	0.260	0.510
Directional	0.010	0.024	0.050	0.100	0.250	0.501
Skovgaard, 2001	0.010	0.025	0.050	0.101	0.251	0.501
<hr/>						
Nominal	0.750	0.900	0.950	0.975	0.990	
LRT	0.757	0.905	0.952	0.974	0.992	
Directional	0.752	0.902	0.950	0.973	0.992	
Skovgaard, 2001	0.752	0.900	0.950	0.973	0.991	

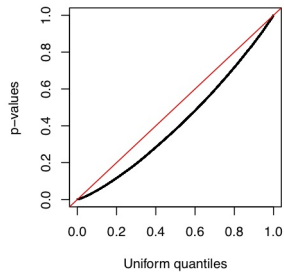
Example: comparison of normal variances



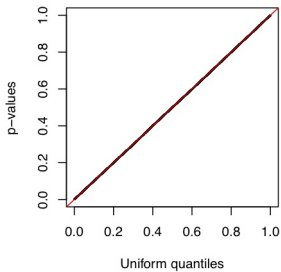
Likelihood ratio statistic	0.0042
Directional	0.0389
Skovgaard, 2001	0.0622
Bartlett's test	0.0136

F-test

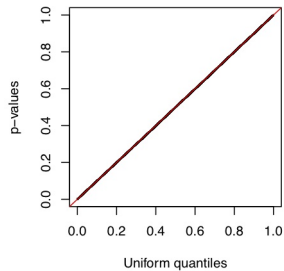
Likelihood ratio statistic



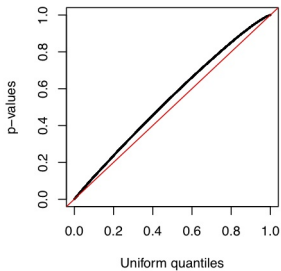
Bartlett's test



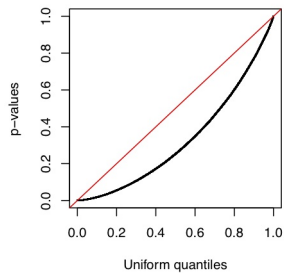
directional



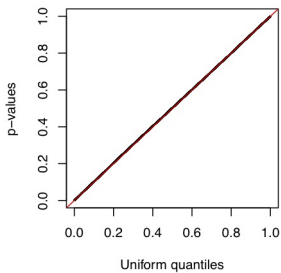
Skovgaard's W^*



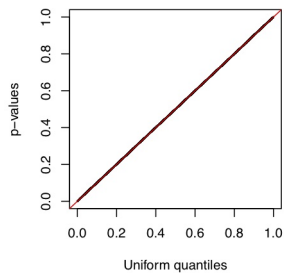
Likelihood ratio statistic



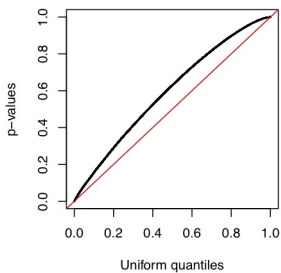
Bartlett's test



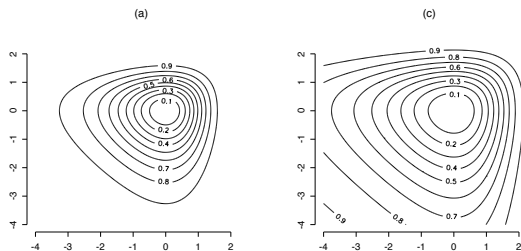
directional



Skovgaard's W^*

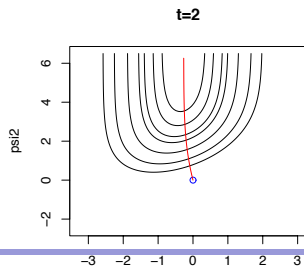
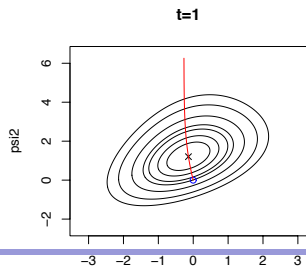
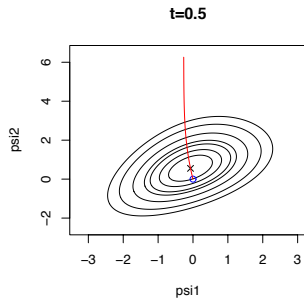
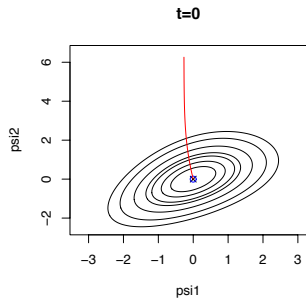


Aside: Bartlett correction of likelihood ratio test



- ▶ $w(\psi) = 2\{\ell(\hat{\theta}) - \ell(\hat{\theta}_\psi)\} \sim \chi_d^2$
- ▶ $E\{w(\psi)\} = d\left\{1 + \frac{B(\psi)}{n}\right\} + O(n^{-2})$
- ▶ $\tilde{w}(\psi) = \frac{w(\psi)}{1 + B(\psi)/n} \sim \chi_d^2 \quad O(n^{-3})$

Aside: compare directional test



Other examples

- ▶ binary regression
- ▶ equality of exponential rates
- ▶ covariance selection models

- ▶ comparison of normal variances
- ▶ symmetry restrictions in contingency tables

- ▶ ...

Binary regression

- ▶ binary response
 - presence/absence of bacteria MASS, §10.4
- ▶ covariate of interest
 - week: factor variable with levels 0, 2, 4, 6, 11
- ▶ nuisance parameter
 - subject effects: factor with 50 levels
- ▶ 108 observations
 - 24 nuisance parameters, 4 parameters of interest

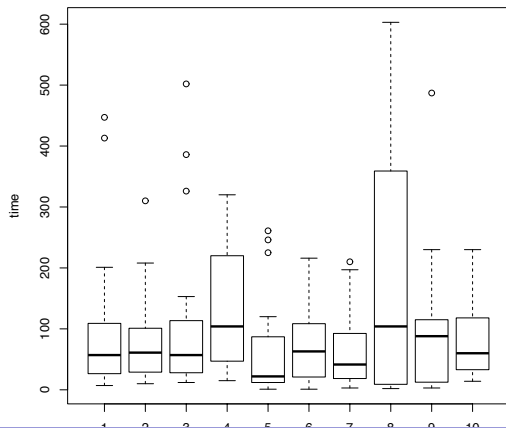
Likelihood ratio statistic 0.0005

Directional **0.0053**

Skovgaard, 2001 0.0043

Equality of exponential rates

- ▶ $y_{ij} \sim \exp(\theta_j)$, $j = 1, \dots, n_i; i = 1, \dots, g$
- ▶ $H_0 : \theta_1 = \dots = \theta_g$
- ▶ Example: air-conditioning data, Cox & Snell (1981)



... exponential rates

- ▶ $p\text{-value} = \frac{\int_1^\infty t^{d-1} h(t) dt}{\int_0^\infty t^{d-1} h(t) dt}$
- ▶ $h(t) \propto \prod_{i=1}^g \left\{ 1 - \frac{t(\bar{y} - \bar{y}_i)}{\bar{y}} \right\}^{n_i-1}$

Nominal	0.010	0.025	0.050	0.100	0.250	0.500
LRT	0.013	0.031	0.059	0.115	0.275	0.529
Directional	0.010	0.025	0.049	0.099	0.248	0.500
Skovgaard, 2001	0.008	0.020	0.039	0.080	0.210	0.441
Nominal	0.750	0.900	0.950	0.975	0.990	
LRT	0.770	0.910	0.955	0.977	0.990	
Directional	0.750	0.899	0.949	0.975	0.989	
Skovgaard, 2001	0.688	0.855	0.918	0.953	0.977	

Conclusion

- ▶ new way to assess vector parameters
- ▶ incorporates information in the direction of departure
- ▶ easy to compute: two model fits, plus 1-d numerical integration
- ▶ accurate conditionally, by construction, and unconditionally (simulations)
- ▶ can be used in models of practical interest – discrete and continuous data
- ▶ exponential family model not necessary – easily generalized using approximate exponential family model

Thank you!



References

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