Composite Likelihood

Nancy Reid

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with Harry Joe, Cristiano Varin and thanks to Don Fraser, Grace Yi, Ximing Xu

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Terminology

- ▶ Model $Y \sim f(y; \theta), \quad y \in \mathbb{R}^m, \quad \theta \in \mathbb{R}^p$
- Events A_1, \ldots, A_k ; "sub-densities" $f(y \in A_k; \theta)$
- Composite log-likelihood

$$\boldsymbol{c\ell}(\theta;\boldsymbol{y}) = \sum_{k=1}^{K} w_k \log f(\boldsymbol{y} \in \boldsymbol{A}_k; \theta) = \sum_{i=1}^{K} w_k \, \ell(\theta; \, \boldsymbol{y} \in \boldsymbol{A}_k)$$

- *w_k* weights to be determined
- composite likelihood is a type of:
 - pseudo-likelihood (spatial modelling);
 - quasi-likelihood (econometrics);
 - limited information method (psychometrics)

► ...

Examples of $c\ell(\theta)$

 $\sum_{r=1}^{m} w_r \log f_1(y_r; \theta) \quad \text{Independence}$ $\sum \sum w_{rs} \log f_2(y_r, y_s; \theta)$ Pairwise $\sum_{r=1}^{m} w_r \log f(y_r \mid y_{(-r)}; \theta) \quad \text{Conditional}$ $\sum \sum w_{rs} \log f(y_r \mid y_s; \theta)$ All pairs conditional $\sum_{r=1}^{m} w_r \log f(y_r \mid y_{r-1}; \theta) \quad \text{Time series}$ $\sum_{r=1}^{\infty} w_r \log f(y_r \mid \text{`neighbours' of } y_r; \theta) \quad \text{Spatial}$ likelihood of (small) blocks of observations; pretend blocks indep. likelihood of pairwise differences your favourite fix here ...

Inference

- Sample y_1, \ldots, y_n independent
- Composite log-likelihood $\sum_{i=1}^{n} c\ell(\theta; y_i)$; maximized at $\hat{\theta}_{CL}$

• As $n \longrightarrow \infty$:

$$\sqrt{n}(\hat{\theta}_{CL}-\theta) \xrightarrow{\mathcal{L}} N\{0, \mathbf{G}^{-1}(\theta)\},$$

• Godambe information $G(\theta) = H(\theta)J^{-1}(\theta)H(\theta)$

•
$$H(\theta) = \mathsf{E}\left\{-\frac{\partial^2 c\ell(\theta; Y_i)}{\partial \theta \partial \theta^T}\right\}, \quad J(\theta) = \mathsf{var}\left\{\frac{\partial c\ell(\theta; Y_i)}{\partial \theta}\right\}$$

... inference

- Sample y_1, \ldots, y_n independent
- Composite log-likelihood $c\ell^n(\theta) = \sum_{i=1}^n c\ell(\theta; y_i);$
- ► CL log-likelihood ratio $w_{CL}(\theta) = 2\{c\ell^n(\hat{\theta}_{CL}) c\ell^n(\theta)\}$

• As
$$n \longrightarrow \infty$$
:

$$W_{CL}(\theta) \xrightarrow{\mathcal{L}} \sum_{j=1}^{p} \lambda_j \chi_{1j}^2$$

• λ_j eigenvalues of $J^{-1}(\theta)H(\theta)$

What do we know?

- ▶ $\hat{\theta}_{CL}$ not fully efficient, unless $G(\theta) = H(\theta)J^{-1}(\theta)H(\theta) = i(\theta)$
- $c\ell(\theta)$ is not a log-likelihood function



- efficiency of $\hat{\theta}_{CL}$ can be pretty high, in many applications
- $w_{CL}(\theta)$ can be re-scaled to $\dot{\sim} \chi_{\rho}^2$

Chandler & Bate 07, Salvan et al. 11

► a little about asymptotics as m → ∞, n fixed or increasing slowly

... what do we know?

- ► careful choice of weights can improve efficiency of $\hat{\theta}_{CL}$ in special cases
- weights can be used to incorporate sampling information, including missing data

Yi 12, Molenberghs 12, Briollais & Choi 12

composite likelihood can be used for model selection

$$AIC_{CL} = -2c\ell^{n}(\hat{\theta}_{CL}) + 2 \operatorname{tr}\{J(\hat{\theta})H^{-1}(\hat{\theta})\}$$

$$BIC_{CL} = -2c\ell^{n}(\hat{\theta}_{CL}) + \log(n)\operatorname{tr}\{J(\hat{\theta})H^{-1}(\hat{\theta})\}$$

- and prediction
- combination of full likelihood for mean parameters and CL for covariance parameters works well in some settings

What don't we know?

- Design
 - marginal vs. conditional
 - choice of weights
 - down-weighting 'distant' observations
 - choosing blocks and block sizes
- Uncertainty estimation
 - $\hat{J}(\hat{\theta}_{CL}) = \hat{var}\{\partial c\ell(\theta)/\partial\theta\}$ need replication; need lots of replication
 - perhaps estimate G(\(\heta_{CL}\)) or var(\(\heta_{CL}\)) directly bootstrap, jackknife
 - or estimate using ideas from higher-order asymptotic approximations
 Fraser 12
 - or try to find some orthogonal components
 Lindsay 12

... what don't we know?

- Identifiability (1): does there exist a model compatible with a set of marginal or conditional densities?
- Identifiability (2): what if different components are estimating different parameters?
- Robustness: CL uses 'low-dimensional' information: is this a type of robustness?
 - find a class of models with same low-d marginals Xu 12
 - classical perturbation of starting model (using copulas?)
 Joe 12
 - random effects models might be amenable to theoretical analysis
 Jordan 12
- asymptotic theory for large *m* (long vectors of responses), small *n*
- relationship to GEE

Some surprises

Y ~ N(μ,Σ) − μ̂_{CL} = μ̂, Σ̂_{CL} = Σ̂ (marginal or conditional (pairwise or full))

•
$$Y \sim \mathcal{N}(\mu \underline{1}, \sigma^2 R), \quad R = \begin{pmatrix} 1 & \rho & \dots & \rho \\ \rho & 1 & \dots & \rho \\ \vdots & \ddots & \ddots & \vdots \\ \rho & \dots & \rho & 1 \end{pmatrix}$$

• $\hat{\theta}_{CL} = \hat{\theta}, \quad G(\theta) = i(\theta), G(\theta) = H(\theta)J^{-1}(\theta)H(\theta)$

•
$$H(\theta) = \text{var}(\text{Score}), J = E(\nabla_{\theta}\text{Score}), H \neq J,$$

•
$$Y \sim (0, R)$$
: $\hat{\rho}_{CL} \neq \hat{\rho}$; a.var $(\hat{\rho}_{CL}) > a.var(\hat{\rho})$

- efficiency improvement when nuisance parameter is unknown
 Mardia et al 08; Xu 12
- CL can be fully efficient, even if $H(\theta) \neq J(\theta)$

... some surprises

- Godambe information G(θ) can decrease as more component CLs are added
- pairwise CL can be less efficient than independence CL
- this can't always be fixed by weighting

parameter constraints can be important

Example: binary vector Y,

$$P(Y_j = y_j, Y_k = y_k) \propto \frac{\exp(\beta y_j + \beta y_k + \theta_{jk} y_j y_k)}{\{1 + \exp(\beta y_j + \beta y_k + \theta_{jk} y_j y_j y_k)\}}$$

- this model is inconsistent
- parameters may not be identifiable in the CL, even if they are in the full likelihood
 Yi, 12
- CL may help get rid of nuisance parameters (e.g. by conditioning)
 Hjort and Varin, 07

Some (more) interesting applications

- spatial data and space-time data
 - conditional approaches seem more natural
 - condition on neighbours (in space); some small number of lags (in time)
 - some form of blockwise components often proposed Stein et al, 04; Caragea and Smith, 07
 - fMRI time series
 Kang et al 12
 - air pollution and health effects
 - computer experiments: Gaussian process models
 Xi 12
- spatially correlated extremes
 - joint tail probability known
 - joint density requires combinatorial effort (partial derivatives)
 - composite likelihood based on joint distribution of pairs, triples seems to work well

Davison et al 12; Genton et al 12

Bai et al 12

... applications

- time series a case of large m, fixed n
 - need new arguments re consistency, asymptotic normality
 - consecutive pairs: consistent, not asy. normal
 - AR(1): consecutive pairs fully efficient; all pairs terrible (consistent, highly variable)
 - MA(1): consecutive pairs terrible

Davis and Yau 11

- genetics: estimation of recombination rate
 - somewhat similar to time series
 - but correlation may not decrease with increasing length
 - suggesting all possible pairs may be inconsistent
 - joint blocks of short sequences seems preferable
- linkage disequilibrium
- family based sampling

Larribe and Fearnhead 11; Choi and Briollais 12

... applications

Gaussian graphical models

Gao and Massam 12

- symmetry constraints have a natural formulation in terms of elements of concentration matrix
- conditional distribution of $y_j \mid y_{(-j)}$
- multivariate binary data for multi-neuron spike trains

Amari 12

CL as a working likelihood in 'maximization by parts'

Bellio 12

- latent variable models in psychometrics
 Maydeu-Olivares 12
- many linear and generalized linear models with random effects
- multivariate survival data

► ...

Some dichotomies

- conditional vs marginal
- pairwise vs everything else
- unstructured vs time series/spatial
- weighted vs unweighted
- "it works" vs "why does it work?" vs "when will it not work"
- ▶ ...