

Likelihood and Asymptotic Theory for Statistical Inference

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London Taught Course Centre

for PhD students in the mathematical sciences

Last week

1. Likelihood – definition, examples, direct inference
2. Derived quantities – score, mle, Fisher information, Bartlett identities
3. Inference from derived quantities – consistency of mle, asymptotic normality
4. Inference via pivotals – standardized score function, standardized mle, likelihood ratio, likelihood root
5. Nuisance parameters and parameters of interest; invariance under parameter transformation
6. Asymptotics for posteriors

This week

1. Bayesian approximation
 - 1.1 careful statement of asymptotic normality
 - 1.2 Laplace approximation to posterior density and cumulative distribution function
 - 1.3 Laplace approximation to marginal posterior density and cdf
 - 1.4 relation to modified profile likelihood

2. Frequentist inference with nuisance parameters
 - 2.1 first order summaries; difficulties with profile likelihood
 - 2.2 marginal and conditional likelihood
 - 2.3 exponential families
 - 2.4 transformation families
 - 2.5 adjustments to profile likelihood

3. Notation

Posterior is asymptotically normal

$$\pi(\theta | \mathbf{y}) \sim N\{\hat{\theta}, j^{-1}(\hat{\theta})\} \quad \theta \in \mathbb{R}, \mathbf{y} = (y_1, \dots, y_n)$$

careful statement

... posterior is asymptotically normal

$$\pi(\theta | \mathbf{y}) \sim N\{\hat{\theta}, j^{-1}(\hat{\theta})\} \quad \theta \in \mathbb{R}, \mathbf{y} = (y_1, \dots, y_n)$$

equivalently $\ell_{\pi}(\theta) =$

... posterior is asymptotically normal

In fact,

If $\pi(\theta) > 0$ and $\pi'(\theta)$ is continuous in a neighbourhood of θ_0 , there exist constants D and n_y s.t.

$$|F_n(\xi) - \Phi(\xi)| < Dn^{-1/2}, \quad \text{for all } n > n_y,$$

on an almost-sure set with respect to $f(y; \theta_0)$, where $y = (y_1, \dots, y_n)$ is a sample from $f(y; \theta_0)$, and θ_0 is an observation from the prior density $\pi(\theta)$.

$$F_n(\xi) = \Pr\{(\theta - \hat{\theta})j^{1/2}(\hat{\theta}) \leq \xi \mid y\}$$

Johnson (1970); Datta & Mukerjee (2004)

Laplace approximation

$$\pi(\theta | \mathbf{y}) \doteq \frac{1}{(2\pi)^{1/2}} |j(\hat{\theta})|^{+1/2} \exp\{\ell(\theta; \mathbf{y}) - \ell(\hat{\theta}; \mathbf{y})\} \frac{\pi(\theta)}{\pi(\hat{\theta})}$$

$$\pi(\theta | \mathbf{y}) = \frac{1}{(2\pi)^{1/2}} |j(\hat{\theta})|^{+1/2} \exp\{\ell(\theta; \mathbf{y}) - \ell(\hat{\theta}; \mathbf{y})\} \frac{\pi(\theta)}{\pi(\hat{\theta})} \{1 + O_p(n^{-1})\}$$

$$\mathbf{y} = (y_1, \dots, y_n), \quad \theta \in \mathbb{R}^1$$

$$\pi(\theta | \mathbf{y}) = \frac{1}{(2\pi)^{1/2}} |j_\pi(\hat{\theta}_\pi)|^{+1/2} \exp\{\ell_\pi(\theta; \mathbf{y}) - \ell_\pi(\hat{\theta}_\pi; \mathbf{y})\} \{1 + O_p(n^{-1})\}$$

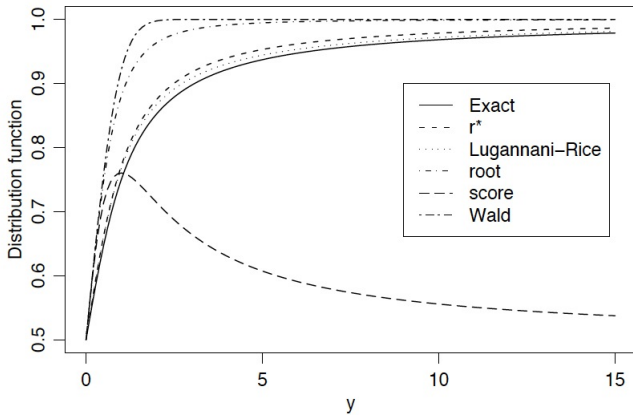
Posterior cdf

$$\int_{-\infty}^{\theta} \pi(\vartheta | y) d\vartheta \doteq \int_{-\infty}^{\theta} \frac{1}{(2\pi)^{1/2}} e^{\ell(\vartheta; y) - \ell(\hat{\vartheta}; y)} |j(\hat{\vartheta})|^{1/2} \frac{\pi(\vartheta)}{\pi(\hat{\vartheta})} d\vartheta$$

Posterior cdf

$$\int_{-\infty}^{\theta} \pi(\vartheta | y) d\vartheta \doteq \int_{-\infty}^{\theta} \frac{1}{(2\pi)^{1/2}} e^{\ell(\vartheta; y) - \ell(\hat{\vartheta}; y)} |j(\hat{\vartheta})|^{1/2} \frac{\pi(\vartheta)}{\pi(\hat{\vartheta})} d\vartheta$$

SM, §11.3



BDR, Ch.3, Cauchy with flat prior

Nuisance parameters

$$\mathbf{y} = (y_1, \dots, y_n) \sim f(\mathbf{y}; \theta), \quad \theta = (\psi, \lambda)$$

$$\begin{aligned} \pi_m(\psi | \mathbf{y}) &= \int \pi(\psi, \lambda | \mathbf{y}) d\lambda \\ &= \frac{\int \exp\{\ell(\psi, \lambda; \mathbf{y})\} \pi(\psi, \lambda) d\lambda}{\int \int \exp\{\ell(\psi, \lambda; \mathbf{y})\} \pi(\psi, \lambda) d\psi d\lambda} \end{aligned}$$

... nuisance parameters

$$y = (y_1, \dots, y_n) \sim f(y; \theta), \quad \theta = (\psi, \lambda)$$

$$\begin{aligned}\pi_m(\psi | y) &= \int \pi(\psi, \lambda | y) d\lambda \\ &= \frac{\int \exp\{\ell(\psi, \lambda; y)\pi(\psi, \lambda) d\lambda}{\int \int \exp\{\ell(\psi, \lambda; y)\pi(\psi, \lambda) d\psi d\lambda}\end{aligned}$$

$$|j(\hat{\theta})| = |j^{\psi\psi}(\hat{\theta})| |j_{\lambda\lambda}(\hat{\theta})|$$

Posterior marginal cdf, $d = 1$

$$\begin{aligned}\Pi_m(\psi | \mathbf{y}) &= \int_{-\infty}^{\psi} \pi_m(\xi | \mathbf{y}) d\xi \\ &\doteq \int_{-\infty}^{\psi} \frac{1}{(2\pi)^{1/2}} e^{\ell_P(\xi) - \ell_P(\hat{\xi})} j_P^{1/2}(\hat{\xi}) \frac{\pi(\xi, \hat{\lambda}_\xi)}{\pi(\hat{\xi}, \hat{\lambda})} \frac{|j_{\lambda\lambda}(\hat{\xi}, \hat{\lambda})|^{1/2}}{|j_{\lambda\lambda}(\xi, \hat{\lambda}_\xi)|^{1/2}} d\xi\end{aligned}$$

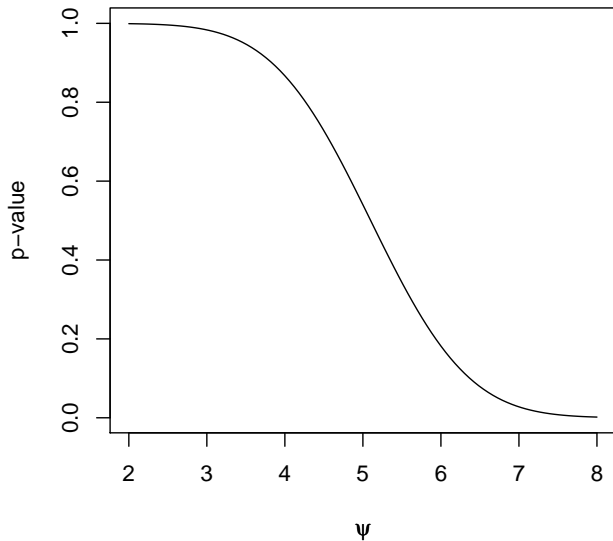
... posterior marginal cdf, $d = 1$

$$\Pi_m(\psi | y) \doteq \Phi(r_B^*) = \Phi\left\{r + \frac{1}{r} \log\left(\frac{q_B}{r}\right)\right\}$$

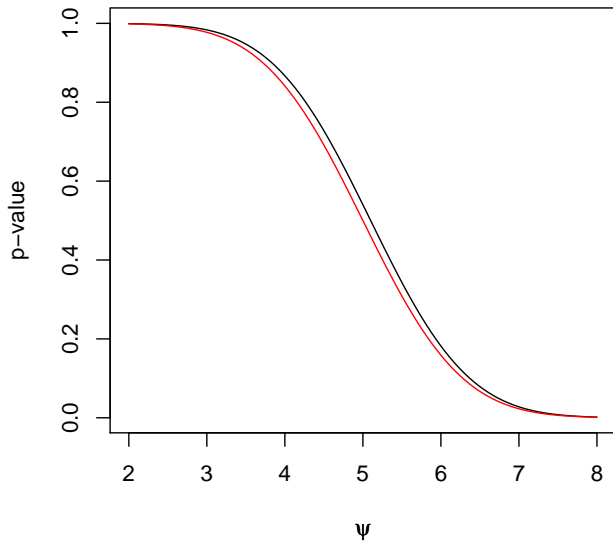
$$r = r(\psi) =$$

$$q_B = q_B(\psi) =$$

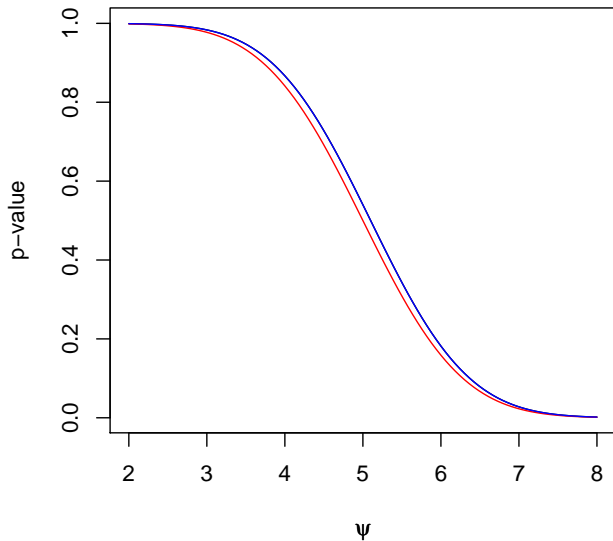
normal circle, k=2



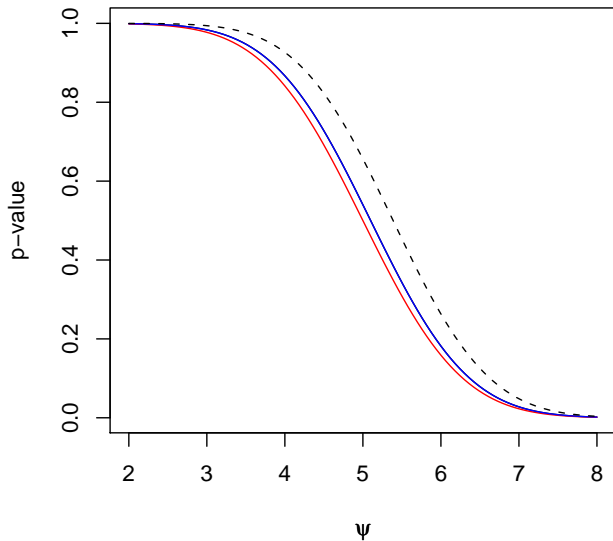
normal circle, k=2



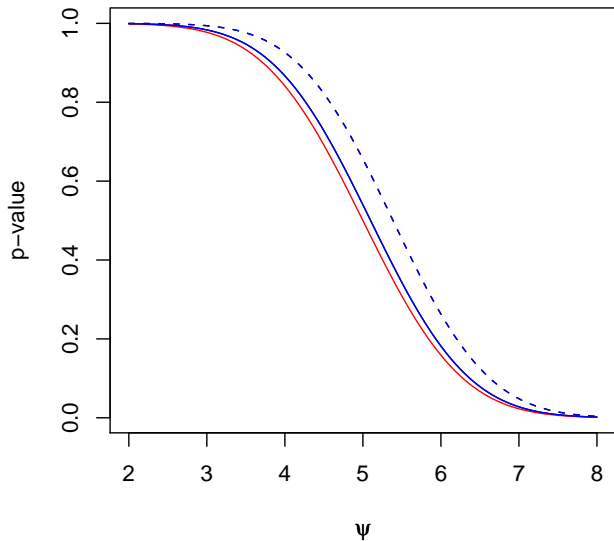
normal circle, k=2



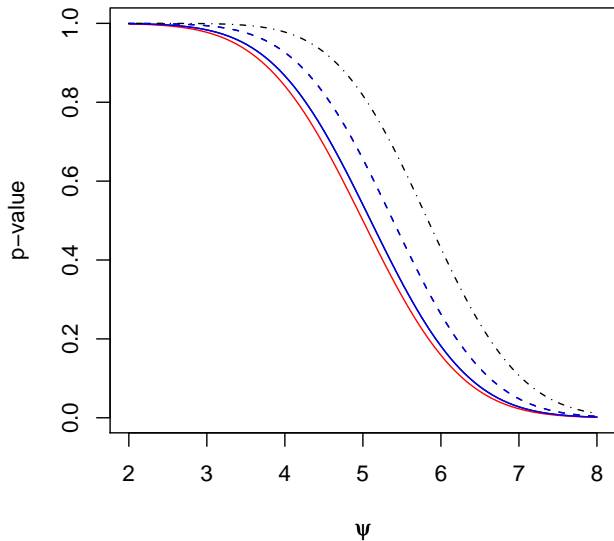
normal circle, $k = 2, 5, 10$



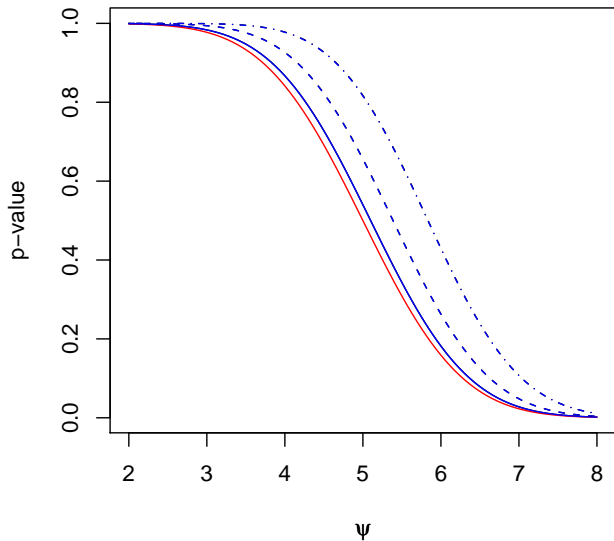
normal circle, $k = 2, 5, 10$



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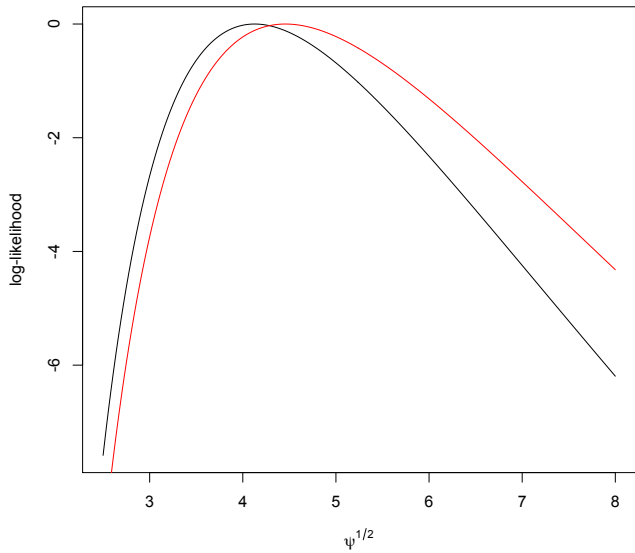
Posterior marginal and adjusted log-likelihoods

$$\pi_m(\psi | \mathbf{y}) \doteq \frac{1}{(2\pi)^{d/2}} e^{\ell_P(\xi) - \ell_P(\hat{\xi})} j_P^{1/2}(\hat{\xi}) \frac{\pi(\xi, \hat{\lambda}_\xi)}{\pi(\hat{\xi}, \hat{\lambda})} \frac{|j_{\lambda\lambda}(\hat{\xi}, \hat{\lambda})|^{1/2}}{|j_{\lambda\lambda}(\xi, \hat{\lambda}_\xi)|^{1/2}}$$

$$\Pi_m(\psi | \mathbf{y}) =$$

Frequentist inference, nuisance parameters

- ▶ first-order pivotal quantities
- ▶ $r_u(\psi) = \ell'_P(\psi)j_P(\hat{\psi})^{1/2} \sim N(0, 1),$
- ▶ $r_e(\psi) = (\hat{\psi} - \psi)j_P(\hat{\psi})^{1/2} \sim N(0, 1),$
- ▶ $r(\psi) = \text{sign}(\hat{\psi} - \psi)2\{\ell_P(\hat{\psi}) - \ell_P(\psi)\} \sim N(0, 1)$
- ▶ all based on treating profile log-likelihood as a one-parameter log-likelihood
- ▶ example $y = X\beta + \epsilon, \quad \epsilon \sim N(0, \psi)$
- ▶ $\hat{\psi} = (y - X\hat{\beta})^T(y - X\hat{\beta})/n$



Eliminating nuisance parameters

- ▶ by using **marginal** density

- ▶ $f(y; \psi, \lambda) \propto f_m(t_1; \psi) f_c(t_2 | t_1; \psi, \lambda)$

- ▶ Example

$$N(X\beta, \sigma^2 I) : f(y; \beta, \sigma^2) \propto f_m(RSS; \sigma^2) f_c(\hat{\beta} | RSS; \beta, \sigma^2)$$

- ▶ by using **conditional** density

- ▶ $f(y; \psi, \lambda) \propto f_c(t_1 | t_2; \psi) f_m(t_2; \psi, \lambda)$

- ▶ Example

$$N(X\beta, \sigma^2 I) : f(y; \beta, \sigma^2) \propto f_c(RSS | \hat{\beta}; \sigma^2) f_m(\hat{\beta}; \beta, \sigma^2)$$

Linear exponential families

- ▶ **conditional density** free of nuisance parameter
- ▶ $f(y_i; \psi, \lambda) = \exp\{\psi^T s(y_i) + \lambda^T t(y_i) - k(\psi, \lambda)\} h(y_i)$
- ▶ $f(y; \psi, \lambda) =$

$s =$

$t =$

- ▶ $f(s, t; \psi, \lambda) =$

- ▶ $f(s | t; \psi) =$

Saddlepoint approximation in linear exponential families

- ▶ no nuisance parameters $f(y_i; \theta) = \exp\{\theta^T \mathbf{s}(y_i) - k(\theta)\} h(y_i)$
- ▶ $f(\mathbf{s}; \theta) = \exp\{\theta^T \mathbf{s} - nk(\theta)\} \tilde{h}(\mathbf{s})$
- ▶ $\ell(\theta; \mathbf{s}) = \theta^T \mathbf{s} - nk(\theta)$
- ▶ $f(\mathbf{s}; \theta) \doteq$

- ▶ $f(\hat{\theta}; \theta) \doteq$

Saddlepoint approximation to conditional density

▶ $f(y_i; \psi, \lambda) = \exp\{\psi^T s(y_i) + \lambda^T t(y_i) - k(\psi, \lambda)\} h(y_i)$

▶ $f(s | t; \psi) =$

▶ $f(\hat{\psi} | t; \psi) \doteq c |j_P(\hat{\psi})|^{1/2} e^{\ell_P(\psi) - \ell_P(\hat{\psi})}$

SM §12.3

Approximating distribution function

- ▶ $f(\hat{\theta}; \theta) \doteq c |j(\hat{\theta})|^{1/2} \exp\{\ell(\theta; \hat{\theta}) - \ell(\hat{\theta}; \hat{\theta})\}$

- ▶ $\int_{-\infty}^{\hat{\theta}} f(\hat{\vartheta}; \theta) d\hat{\vartheta} \doteq$

Summary

- ▶ No nuisance parameters
 - ▶ Bayesian p -value $\Phi(r_B^*)$
 - ▶ $r_B^* = r + \frac{1}{r} \log \frac{q_B}{r}$

 - ▶ Exponential family p -value $\Phi(r^*)$
 - ▶ $r^* = r + \frac{1}{r} \log \frac{q}{r}$

- ▶ Nuisance parameters
 - ▶ Bayesian p -value $\Phi(r_B^*)$
 - ▶ $r_B^* = r + \frac{1}{r} \log \frac{q_B}{r}$

 - ▶ Exponential family p -value $\Phi(r^*)$
 - ▶ $r^* = r + \frac{1}{r} \log \frac{q}{r}$

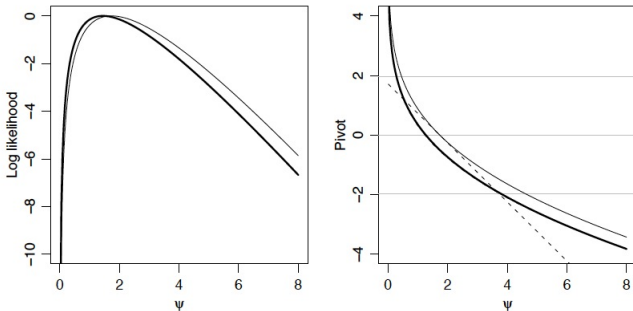


Figure 2.3: Inference for shape parameter ψ of gamma sample of size $n = 5$. Left: profile log likelihood ℓ_p (solid) and the log likelihood from the conditional density of u given v (heavy). Right: likelihood root $r(\psi)$ (solid), Wald pivot $t(\psi)$ (dashes), modified likelihood root $r^*(\psi)$ (heavy), and exact pivot overlying $r^*(\psi)$. The horizontal lines are at $0, \pm 1.96$.

