## STA 302 / 1001 H - Fall 2003 Test 1

October 6, 2003

LAST NAME: <b>SOLUTIONS</b>	FIRST NAM	FIRST NAME:				
STUDENT NUMBER:						
ENROLLED IN: (circle one)	STA 302	STA 1001				

## INSTRUCTIONS:

- Time: 50 minutes
- Aids allowed: calculator.
- A table of values from the t distribution is on the last page.
- Total points: 30

## Some formulae:

$$b_{1} = \frac{\sum (X_{i} - \overline{X})(Y_{i} - \overline{Y})}{\sum (X_{i} - \overline{X})^{2}}$$

$$Var(b_{1}) = \frac{\sigma^{2}}{\sum (X_{i} - \overline{X})^{2}}$$

$$Var(b_{0}) = \sigma^{2} \left(\frac{1}{n} + \frac{\overline{X}^{2}}{\sum (X_{i} - \overline{X})^{2}}\right)$$

$$SSTO = \sum (Y_{i} - \overline{Y})^{2}$$

$$SSE = \sum (Y_{i} - \hat{Y}_{i})^{2}$$

$$SSR = b_{1}^{2} \sum (X_{i} - \overline{X})^{2} = \sum (\hat{Y}_{i} - \overline{Y})^{2}$$

$$\sigma^{2}\{\hat{Y}^{*}\} = Var(\hat{Y}^{*}) = \sigma^{2} \left(\frac{1}{n} + \frac{(X^{*} - \overline{X})^{2}}{\sum (X_{i} - \overline{X})^{2}}\right) \quad \sigma^{2}\{pred\} = Var(Y^{*} - \hat{Y}^{*}) = \sigma^{2} \left(1 + \frac{1}{n} + \frac{(X^{*} - \overline{X})^{2}}{\sum (X_{i} - \overline{X})^{2}}\right)$$

$$1 \text{ (a) (b)} \qquad 1 \text{ (c) (d) (e)} \qquad 2 \text{ (a) (b) (c)} \qquad 3$$

1. (a) (4 points) State the simple linear regression model for dependent variable Y and independent variable X and the Gauss-Markov conditions.

Model: 
$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

 $Gauss-Markov\ conditions:$ 

$$E(\epsilon_i) = 0$$

$$Var(\epsilon_i) = \sigma^2$$

 $\epsilon_i$ 's are uncorrelated

(The GM conditions can also be stated in terms of the Yi's.)

(b) (5 points) The least squares estimate of the Y intercept for the model in (a) is  $b_0$  as given on the first page. Under the model you specified in (a), derive the formula for the variance of  $b_0$ . You may **not** asssume that any of the formulae for variance or covariance are known.

Sketch of solution:

First show that  $b_0$  is a linear combination of the  $Y_i$ 's, i.e.

$$b_0 = \sum_{i=1}^{n} \left( \frac{1}{n} - \frac{\overline{X}(X_i - \overline{X})}{\sum (X_i - \overline{X})^2} \right) Y_i$$

Since the  $Y_i$ 's are uncorrelated,  $Var(\sum k_iY_i) = \sum k_i^2 Var(Y_i) = \sum k_i^2 \sigma^2$ Squaring the  $k_i$ 's and summing gives the result.

- (c) (3 points) In order to do inference about the slope (such as testing whether or not the slope is 0) we need to make one more assumption about the model in (a). What is the usual assumption and why is it necessary?

  The errors have a normal distribution. We need a distributional assumption for the errors in order to have a sampling distribution for the estimators which is used for critical values and p-values.
- (d) (2 points) An estimate is more precise than another if it has smaller variance. The estimate of the expected value of Y varies with the value of X. At what value of X will there be the most precise estimate of the expected value of Y? Justify your answer.

The variance of the expected value of Y at  $X^*$  is

$$\sigma^2 \left( \frac{1}{n} + \frac{(X^* - \overline{X})^2}{\sum (X_i - \overline{X})^2} \right)$$

This is smallest when  $X^*$  is the average value of the X's.

(e) (2 points) Suppose the dependent variable Y could be measured with less error. Why would this lead to more precise estimation of the intercept of the regression line?

The variance of the intercept is  $\sigma^2\left(\frac{1}{n} + \frac{\overline{X}^2}{\sum (X_i - \overline{X})^2}\right)$ . If Y has less error,  $\sigma^2$  is smaller, so the variance of the intercept is smaller.

2. Some engineers are interested in examining the relationship between load, in pounds, and the corresponding deformation, in inches, on a mild steel bar. Based on the physical properties of the steel, the engineers believe there will be a linear relationship between the natural logarithm of load (logL) and the natural logarithm of deformation (logD). Data were collected for 24 loads ranging from 1000 to 9900 pounds and a regression of logD on logL was run. Some output from SAS is given below. (Some numbers have been purposely removed from the output.)

Descriptive Statistics										
Uncorrected							${\tt Standard}$			
Variable Sum		Mean		SS	S Variance		Deviation			
Intercept	24.00000	1.00	000 24	.00000		0	0			
logL	212.18446	8.84	102 1883	. 83553	0.34	386	0.58639			
logD	174.88129	7.28	672 1304	. 29718	1.30	375	1.14182			
Dependent Veriable, lead										
Dependent Variable: logD										
Analysis of Variance										
_			Sum of		ean _					
Source		DF	Squares	Squ	are F	Value	Pr > F			
Model		1	20.57353	20.57	353	48.09	<.0001			
Error		22	9.41263	0.42	785					
Corrected Tot	al	23	29.98616							
Parameter Estimates										
		Paramete	r Star	ndard						
Variable	DF	Estimat	e I	Error	t Value	Pr >	t			
Intercep	t 1	-6.9727	5 2.0	06066	-3.38	0.	0027			
logL	1	1.6128	8 0.2	23259	6.93	<.	0001			

The following questions (labelled (a) through (d)) relate to the SAS output on the previous page.

- (a) (3 points) Construct a 95% confidence interval for the slope. What can you conclude from this confidence interval about a test of  $H_0: \beta_1 = 0$ ?  $t_{22,0.025} = 2.074$  95% CI for  $\beta_1: 1.61288 \pm 2.074(0.233259) = (1.1305, 2.0953)$  The p-value for a two-sided test of  $H_0: \beta_1 = 0$  is less than 0.05.
- (b) (2 points) What is the value of  $R^2$ ? What does this value mean?  $R^2 = 20.57353/29.98616 = 0.686$  68.6% of the variation in the log of deformation is explained by its linear relationship with the log of load.
- (c) (4 points) What is the predicted value of logD when the load is 9600 pounds? Construct an appropriate 90% interval for the expected value of logD at this load.

The predicted value of logD when the load is 9600 pounds is  $-6.97275 + 1.61288[\log(9600)] = 7.8166$   $t_{22,.05} = 1.717$  90% interval for mean of logD when load is 9600:

$$7.8166 \pm 1.717\sqrt{.42785}\sqrt{\left(\frac{1}{24} + \frac{[log(9600) - 8.84102]^2}{(23)(.34386)}\right)} = (7.5525, 8.0807)$$

- (d) (2 points) Do you trust your prediction in (c)? Explain.

  No. As can be seen from the scatterplot, a linear model is not appropriate.
- 3. (3 points) A simple linear regression is performed to study the relationship between a child's intelligence at age 5 (the dependent variable) and the length of time the child was breastfed as an infant (the independent variable). The value of r² was 0.56 and for the two-sided test of whether the slope was 0, the p-value was 0.013. A newspaper headline reporting on the study said "Study shows that to make a smarter child, mothers should breastfeed longer". Comment on the validity of the headline.

While there is an indication of a linear relationship between intelligence at age 5 and length of time breastfed, the headline claims breastfeeding causes intelligence. This is an observational study so causality can't be claimed.