

STA 257 - Solutions to additional questions
from Practice Problems 10

5.68 (a) $M_{X_1, X_2, X_3}(t, t, t) = E(e^{t(X_1 + X_2 + X_3)}) = M_{X_1 + X_2 + X_3}(t)$

(b) $M_{X_1, X_2, X_3}(t, t, 0) = E(e^{t(X_1 + X_2)}) = M_{X_1 + X_2}(t)$

(c)
$$\frac{\partial^{k_1 + k_2 + k_3} M(t_1, t_2, t_3)}{\partial t_1^{k_1} \partial t_2^{k_2} \partial t_3^{k_3}} \Big|_{t_1 = t_2 = t_3 = 0} = E\left(\frac{\partial^{k_1 + k_2 + k_3} e^{t_1 X_1 + t_2 X_2 + t_3 X_3}}{\partial t_1^{k_1} \partial t_2^{k_2} \partial t_3^{k_3}}\right) \Big|_{t_1 = t_2 = t_3 = 0}$$

$$= E(X_1^{k_1} X_2^{k_2} X_3^{k_3} e^{t_1 X_1 + t_2 X_2 + t_3 X_3}) \Big|_{t_1 = t_2 = t_3 = 0}$$

$$= E(X_1^{k_1} X_2^{k_2} X_3^{k_3})$$

7.40 Let $X_i \sim \chi^2_{(1)}$ and X_1, \dots, X_n be independent.

$$Y = \sum_{i=1}^n X_i \sim \chi^2(n)$$

You can use mgfs to show this:

$$m_Y(t) = \int_0^\infty e^{yt} \frac{1}{2^{n/2} \Gamma(n/2)} y^{n/2 - 1} e^{-y/2} dy = \dots = \left[\frac{1}{2(\frac{1}{2} - t)} \right]^{n/2}$$

$$= [m_{X_i}(t)]^n$$

$$= m_{\sum X_i}(t)$$

$$\bar{X} = Y/n$$

By the Central Limit Theorem,

$\frac{\sqrt{n}(\bar{X} - E(X_i))}{\sqrt{V(X_i)}}$ converges in distribution to a standard normal r.v.

$$= \frac{n}{n} * \left(\quad \quad \right) = \frac{n}{n} \left(\frac{\sqrt{n}(\bar{X} - 1)}{\sqrt{2}} \right) \quad \text{since can show } E X_i = 1, \sqrt{X_i} = 2$$

$$= \frac{Y - n}{\sqrt{2n}}$$

$$\#5/ \pi'(t) = \frac{2qp^2}{(1-qt)^3}, \quad \pi''(t) = \frac{6q^2p^2}{(1-qt)^4}$$

$$\pi'(1) = \frac{2q}{p}, \quad \pi''(1) = \frac{6q^2}{p^2}$$

$$= EX$$

$$VX = \frac{6q^2}{p^2} + \frac{2q}{p} - \left(\frac{2q}{p}\right)^2$$

$$\#6/ \text{Maclaurin series of } \pi(t) \text{ is } \pi(t) = 2C \left(1 + \frac{t^2}{2!} + \frac{t^4}{4!} + \frac{t^6}{6!} + \dots\right)$$

$$(a) P(X=1) = 0, \quad P(X=10) = \frac{2C}{10!}$$

$$(b) \pi'(t) = C(e^t - e^{-t}); \quad \pi'(1) = C(e - e^{-1})$$

$$(c) \pi(1) = 1 \Rightarrow 1 = 2C(e + e^{-1})$$

$$\#7/ (a) m_2(t) = e^{t^2/2}; \quad m_X(t) = e^{\mu t} e^{\sigma^2 t^2/2}$$

$$m_X'(t) = \mu e^{\mu t} + \sigma^2 t e^{\sigma^2 t^2/2}; \quad m_X'(0) = \mu = EX$$

$$m_X''(t) = \mu^2 e^{\mu t} + \sigma^2 e^{\sigma^2 t^2/2} + \sigma^4 t^2 e^{\sigma^2 t^2/2}; \quad m_X''(0) = \mu^2 + \sigma^2$$

$$VX = \mu^2 + \sigma^2 - (\mu^2)$$

$$\#8/ (a) m(t) = \sum_{k=0}^{\infty} \frac{e^{-\lambda} (e^t \lambda)^k}{k!} = e^{-\lambda} e^{\lambda e^t}$$

$$(b) \pi(t) = e^{\lambda t - \lambda}$$

$$m(t) = \pi(e^t) = e^{\lambda e^t - \lambda}$$

#9/ Let \bar{X} be the average of the 75 deposits
By the CLT, $\bar{X} \approx N(638, \frac{(126)^2}{75})$

$$P(\bar{X} > 671) \approx P\left(Z > \frac{671-638}{126/\sqrt{75}}\right) \text{ where } Z \sim N(0,1)$$

$$= 0.0116$$

#10/ (b) Chebyshev's inequality:

$$P(|\bar{X}_{50} - E\bar{X}_{50}| \geq \epsilon) \leq \frac{\sqrt{\bar{X}_{50}}}{\epsilon^2}$$

$$P(\bar{X}_{50} \text{ fails to be between } 98 \text{ and } 102)$$

$$= P(|\bar{X}_{50} - 100| \geq 2)$$

$$\leq \frac{1/2}{2^2}$$

(c) By the CLT, $\bar{X}_{50} \approx N(100, \frac{1}{2})$

$$P(\bar{X}_{50} \text{ fails to be between } 98 \text{ and } 102)$$

$$= 1 - P(98 \leq \bar{X}_{50} \leq 102)$$

$$\approx 1 - \left[\Phi\left(\frac{102-100}{\sqrt{1/2}}\right) - \Phi\left(\frac{98-100}{\sqrt{1/2}}\right) \right]$$

$$= .0046$$

where Φ is the cdf
of the standard
normal distribution.