

1. The method of least squares is used to fit a simple linear regression model $Y = \beta_0 + \beta_1 X + \epsilon$ to n observations (X_i, Y_i) . The values of the X_i 's are not realizations of random variables but are fixed in advance by the researcher. Assume the following: the form of the model is appropriate, the Gauss-Markov conditions hold, and the distribution of the error terms is Normal. In your answers, you may use any of the formulae on the front page.

- (a) (4 marks) Which of the assumptions are needed to fit the model using least squares? How would you assess the necessary assumptions?

For least squares, the only assumption needed is the form of the model.

To check this look at plots of Y versus X and the residuals versus either X or the predicted values. Look for outliers / influential points and curvature.

- (b) (3 marks) What is $E(Y)$? Which of the assumptions did you use to determine your answer?

$$E(Y) = \beta_0 + \beta_1 X$$

Assumptions used: form of model and $E(\epsilon) = 0$

- (c) (2 marks) Suppose the researcher is interested in the relationship between X and Y on a certain range of X 's. She uses the smallest value in the range of X for half of the observations and the largest value in the range of X for the other half of the observations and fits the simple linear regression model to the resulting data. What is the advantage of fixing the X 's to be these values? What is the disadvantage?

Advantage: This choice of X 's will make S_{XX} as large as possible, giving more precise estimates (smaller variance) of the model parameters.

Disadvantage: Won't be able to determine if the form of the relationship really is linear without observations on more values of X .

(Question 1 continued)

(d) (4 marks) Show that the least squares estimator b_1 is an unbiased estimator of β_1 .

$$\begin{aligned} E(b_1) &= E\left(\frac{\sum X_i Y_i - n\bar{X}\bar{Y}}{S_{XX}}\right) \\ &= \frac{1}{S_{XX}} \left(\sum X_i E(Y_i) - n\bar{X} \frac{1}{n} \sum E(Y_i)\right) \\ &= \frac{1}{S_{XX}} \left(\sum X_i (\beta_0 + \beta_1 X_i) - \bar{X} \sum (\beta_0 + \beta_1 X_i)\right) \\ &= \frac{1}{S_{XX}} \left(\beta_0 n\bar{X} + \beta_1 \sum X_i^2 - \bar{X} n\beta_0 - \beta_1 \bar{X} n\bar{X}\right) \\ &= \frac{\beta_1}{S_{XX}} \left(\sum X_i^2 - n\bar{X}^2\right) \\ &= \beta_1 \end{aligned}$$

(e) (2 marks) Are b_0 (the estimator of the intercept) and b_1 (the estimator of the slope) independent? Explain.

No. Their covariance is not 0 (unless $\bar{X} = 0$).

2. CFC-11 atmospheric concentrations in parts per trillion were measured monthly. The following SAS output shows the results of the regression of atmospheric concentration on time (in years) for the period 1977 to 1989. On the next page is SAS output for the regression for the period 1995 to 2004. Some of the output has been removed and, in the first regression, some of the numerical values have been replaced by letters. Answer the questions assuming that the usual regression model assumptions hold.

Before Montreal Protocol (before January 1990)

Descriptive Statistics

Variable	Sum	Mean	Uncorrected SS	Variance	Standard Deviation
Intercept	153.00000	1.00000	153.00000	0	0
time	303466	1983.43464	601906139	14.16876	3.76414
cfc11	30286	197.94771	6199037	1342.05751	36.63410

The REG Procedure
Dependent Variable: cfc11

Analysis of Variance					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	(A)	203119	203119	(B)	<.0001
Error	151	874.00094	5.78809		
Corrected Total	(C)	203993			

Root MSE	(D)	R-Square	0.9957
Dependent Mean	197.94771		

Parameter Estimates					
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	1	-19064	(E)	-185.40	<.0001
time	1	9.71152	0.05184	187.33	<.0001

- (a) (5 marks) Find the values of the numbers that have been replaced by letters:

$$(A) = \underline{\quad 1 \quad}$$

$$(B) = \underline{\quad 203119/5.78809 = 35092 \quad}$$

$$(C) = \underline{\quad 152 \quad}$$

$$(D) = \underline{\quad \sqrt{5.78809} = 2.406 \quad}$$

$$(E) = \underline{\quad -19064 / -185.40 = 102.8 \quad}$$

(Question 2 continued)

After Montreal Protocol (after December 1994)

Descriptive Statistics

Variable	Sum	Mean	Uncorrected SS	Variance	Standard Deviation
Intercept	116.00000	1.00000	116.00000	0	0
time	231985	1999.86710	463939247	7.92451	2.81505
cfc11	30641	264.14741	8096893	27.17849	5.21330

The REG Procedure
Dependent Variable: cfc11

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	3061.55314	3061.55314	5455.70	<.0001
Error	114	63.97289	0.56117		
Corrected Total	115	3125.52602			

Root MSE 0.74911
Dependent Mean 264.14741

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	1	3929.67750	49.62630	79.19	<.0001
time	1	-1.83289	0.02481	-73.86	<.0001

- (b) (2 marks) Calculate a 90% confidence interval for the intercept for the regression for the period 1995 to 2004 (after the Montreal Protocol).

$$t_{114,05} = 1.671 \text{ (approximating with 60 d.f.)}$$

$$\text{Confidence interval: } 3929.6775 \pm 1.671(49.62630) = (3846.8, 4012.6)$$

- (c) (2 marks) For the regression for the time period 1995 to 2004, find R^2 and explain what it measures.

$$R^2 = 3061.55314/3125.52602 = 0.9795$$

Almost 98% of the variation in CFC-11 concentration for this time period is explained by its linear relationship with time.

- (d) (1 mark) Interpret the estimated slope in practical terms.

On average, in this time period CFC-11 is going down 1.83 parts per trillion per year.

(Question 2 continued)

- (e) (4 marks) Carry out an hypothesis test to determine whether the slopes of the lines for the regressions for the two time periods differ. If you do not have all the information you need to completely answer the question, indicate what is missing and give the most complete answer you can.

s.d. of the difference in the slopes is $\sqrt{0.05184^2 + 0.02481^2} = 0.05747$

Testing $H_0 : \beta_{1,before MP} = \beta_{1,after MP}$ versus $H_a : \beta_{1,before MP} \neq \beta_{1,after MP}$

Test statistic: $(9.71152 - (-1.83289))/0.05747 = 200.9$

In order to estimate the p-value, we need to know the appropriate degrees of freedom for the t-distribution of the test statistic. However, 200.9 is far in the tails of all t distributions so the p-value will be very small.

So we have strong evidence that the 2 slopes differ.

- (f) (4 marks) Use one of the fitted models to predict what the atmospheric concentration of CFC-11 on October 1, 2009 was (when `time` = 2009.75) and give a 99% interval for your prediction.

(Use the second model since it is closer in time to 2009.)

On October 1, 2009, $C\hat{F}C-11 = 3929.6675 - 1.83289(2009.75) = 246.0$

$t_{114,0.005} = 2.660$ (estimating with 60 d.f.)

Prediction interval: $246.0 \pm 2.660(0.74911)\sqrt{1 + \frac{1}{116} + \frac{(2009.75 - 1999.8671)^2}{115(7.9245)}} = (243.9, 248.1)$

- (g) (2 marks) Do you feel confident that the actual concentration of CFC-11 measured on October 1, 2009 is in the interval you calculated in part (f)? Why or why not?

No because October 1, 2009 is outside the range of the data and we can't be sure that the linear model is still appropriate after 2004.

- (h) (2 marks) Using only what you know about how the data were collected, does it seem possible that there are any violations in the Gauss-Markov conditions for these regressions? Explain.

Yes since the data are collected over time there are likely non-zero correlations in the ϵ_i 's for observations close together.

3. Golf tournaments take place over a few days. On each day of the tournament one round of golf is played. In this question, we are looking at the relationship between golfers' scores on the first round and their scores on the second round in the 2000 British Open. In golf, low scores are good. Some output from SAS is given below.

The REG Procedure
Dependent Variable: round2

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	29.51923	29.51923	3.10	0.0804
Error	154	1467.22436	9.52743		
Corrected Total	155	1496.74359			

Root MSE	3.08665	R-Square	0.0197
Dependent Mean	72.24359	Adj R-Sq	0.0134
Coeff Var	4.27256		

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	1	62.42741	5.58218	11.18	<.0001
round1	1	0.13449	0.07641	1.76	0.0804

- (a) (2 marks) Is there evidence of a linear relationship between golf scores on the first and second round of the tournament? Explain.

The p-value for the test with null hypothesis that the slope is 0 is 0.0804 so we have only weak evidence of a linear relationship.

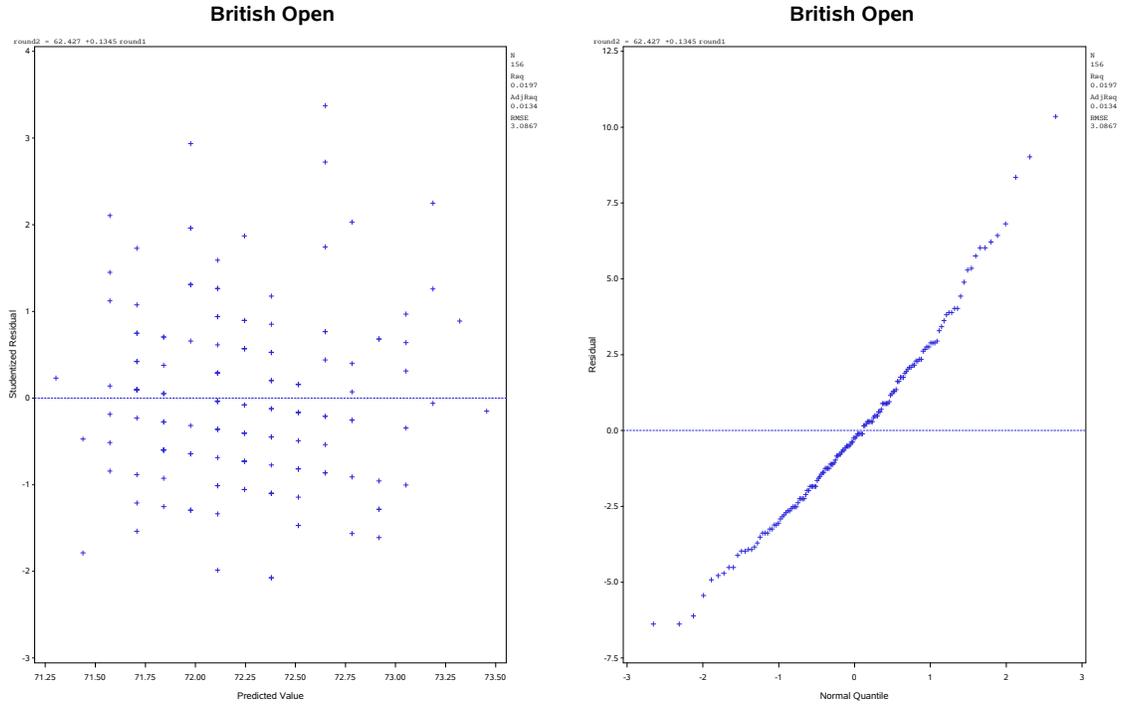
- (b) (3 marks) The lowest score obtained on the first round was 66. Predict the second round score of the golfer who achieved this. Is this surprising? Explain your answer in terms of known facts about simple linear regression.

$$\widehat{\text{round2}} = 62.42741 + 0.13449(66) = 71$$

So we expect the second round score to jump to up 71. It is not surprising to see it go up as this is an example of regression to the mean. In any test/re-test situation we expect people with low scores the first time to have higher scores, on average, the second time.

(Question 3 continued)

- (c) (4 marks) Below are plots of the studentized residuals versus the predicted values, and a normal quantile plot of the residuals. What additional information do you learn from the plots? Be specific.



The first plot looks like random scatter about 0. So there are no problems with: outliers / influential points, curvature, constant variance.

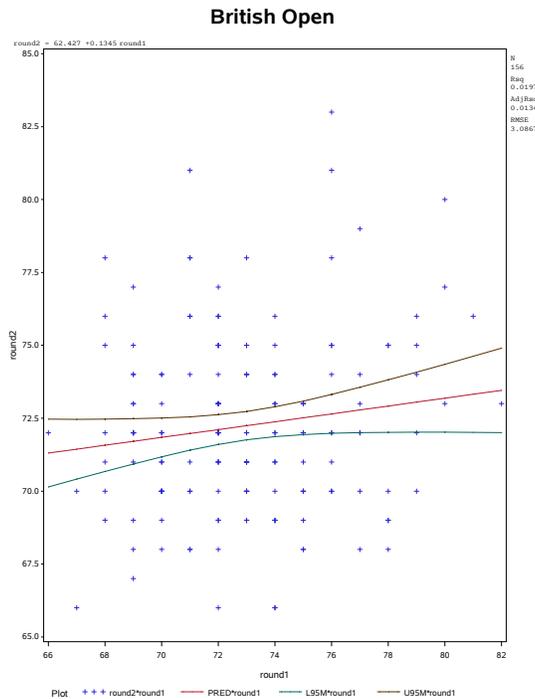
The second plot looks like a straight line so there is no problem with the assumption that the errors are normally distributed.

(Question 3 continued)

- (d) (2 marks) In assignment 1 we considered the relationship between football kickers' field goal percentages one year with the percentage of field goals scored the previous year and found problems with violations of the Gauss-Markov conditions in the initial analysis. In the regression here we are examining the relationship between golf scores on one round with golf scores on the previous round. Do we have a similar problem with violations of the Gauss-Markov assumptions? Why or why not?

No because in this example we don't have multiple observations on the same golfers.

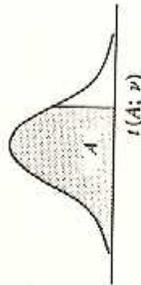
- (e) (2 marks) The plot of the data below includes 95% confidence intervals for the mean score in round 2 given the score in round 1. About 90% of the data points fall outside the confidence limits. Explain how it can occur that so many observations are missed.



These are CIs for $E(Y)$ and only take into account error in the estimation of the regression line and not the fact that individual points vary about the line (the ϵ component of the model). We would expect prediction intervals that also include this other variability to capture more of the points.

TABLE B.2 Percentiles of the *t* Distribution.

Entry is $t(A; \nu)$ where $P\{t(\nu) \leq t(A; \nu)\} = A$



ν	A									
	.60	.70	.80	.85	.90	.95	.975			
1	0.325	0.727	1.376	1.963	3.078	6.314	12.706			
2	0.289	0.617	1.061	1.386	1.886	2.920	4.303			
3	0.277	0.584	0.978	1.250	1.638	2.353	3.182			
4	0.271	0.569	0.941	1.190	1.533	2.132	2.776			
5	0.267	0.559	0.920	1.156	1.476	2.015	2.571			
6	0.265	0.553	0.906	1.134	1.440	1.943	2.447			
7	0.263	0.549	0.896	1.119	1.415	1.895	2.365			
8	0.262	0.546	0.889	1.108	1.397	1.860	2.306			
9	0.261	0.543	0.883	1.100	1.383	1.833	2.262			
10	0.260	0.542	0.879	1.093	1.372	1.812	2.228			
11	0.260	0.540	0.876	1.088	1.363	1.796	2.201			
12	0.259	0.539	0.873	1.083	1.356	1.782	2.179			
13	0.259	0.537	0.870	1.079	1.350	1.771	2.160			
14	0.258	0.537	0.868	1.076	1.345	1.761	2.145			
15	0.258	0.536	0.866	1.074	1.341	1.753	2.131			
16	0.258	0.535	0.865	1.071	1.337	1.746	2.120			
17	0.257	0.534	0.863	1.069	1.333	1.740	2.110			
18	0.257	0.534	0.862	1.067	1.330	1.734	2.101			
19	0.257	0.533	0.861	1.066	1.328	1.729	2.093			
20	0.257	0.533	0.860	1.064	1.325	1.725	2.086			
21	0.257	0.532	0.859	1.063	1.323	1.721	2.080			
22	0.256	0.532	0.858	1.061	1.321	1.717	2.074			
23	0.256	0.532	0.858	1.060	1.319	1.714	2.069			
24	0.256	0.531	0.857	1.059	1.318	1.711	2.064			
25	0.256	0.531	0.856	1.058	1.316	1.708	2.060			
26	0.256	0.531	0.856	1.058	1.315	1.706	2.056			
27	0.256	0.531	0.855	1.057	1.314	1.703	2.052			
28	0.256	0.530	0.855	1.056	1.313	1.701	2.048			
29	0.256	0.530	0.854	1.055	1.311	1.699	2.045			
30	0.256	0.530	0.854	1.055	1.310	1.697	2.042			
40	0.255	0.529	0.851	1.050	1.303	1.684	2.021			
60	0.254	0.527	0.848	1.045	1.296	1.671	2.000			
120	0.254	0.526	0.845	1.041	1.289	1.658	1.980			
∞	0.253	0.524	0.842	1.036	1.282	1.645	1.960			

TABLE B.2 (continued) Percentiles of the *t* Distribution.

ν	A									
	.98	.985	.99	.9925	.995	.9975	.9995			
1	15.895	21.205	31.821	42.434	63.657	127.322	636.590			
2	4.849	5.643	6.965	8.073	9.925	14.089	31.598			
3	3.482	3.896	4.541	5.047	5.841	7.453	12.924			
4	2.999	3.298	3.747	4.088	4.604	5.598	8.610			
5	2.757	3.003	3.365	3.634	4.032	4.773	6.859			
6	2.612	2.829	3.143	3.372	3.707	4.317	5.959			
7	2.517	2.715	2.998	3.203	3.499	4.029	5.408			
8	2.449	2.634	2.896	3.085	3.355	3.833	5.041			
9	2.398	2.574	2.821	2.998	3.250	3.690	4.781			
10	2.359	2.527	2.764	2.932	3.169	3.581	4.587			
11	2.328	2.491	2.718	2.879	3.106	3.497	4.437			
12	2.303	2.461	2.681	2.836	3.055	3.428	4.318			
13	2.282	2.436	2.650	2.801	3.012	3.372	4.221			
14	2.264	2.415	2.624	2.771	2.977	3.326	4.140			
15	2.249	2.397	2.602	2.746	2.947	3.286	4.073			
16	2.235	2.382	2.583	2.724	2.921	3.252	4.015			
17	2.224	2.368	2.567	2.706	2.898	3.222	3.965			
18	2.214	2.356	2.552	2.689	2.878	3.197	3.922			
19	2.205	2.346	2.539	2.674	2.861	3.174	3.883			
20	2.197	2.336	2.528	2.661	2.845	3.153	3.849			
21	2.189	2.328	2.518	2.649	2.831	3.135	3.819			
22	2.183	2.320	2.508	2.639	2.819	3.119	3.792			
23	2.177	2.313	2.500	2.629	2.807	3.104	3.768			
24	2.172	2.307	2.492	2.620	2.797	3.091	3.745			
25	2.167	2.301	2.485	2.612	2.787	3.078	3.725			
26	2.162	2.296	2.479	2.605	2.779	3.067	3.707			
27	2.158	2.291	2.473	2.598	2.771	3.057	3.690			
28	2.154	2.286	2.467	2.592	2.763	3.047	3.674			
29	2.150	2.282	2.462	2.586	2.756	3.038	3.659			
30	2.147	2.278	2.457	2.581	2.750	3.030	3.646			
40	2.123	2.250	2.423	2.542	2.704	2.971	3.551			
60	2.099	2.223	2.390	2.504	2.660	2.915	3.460			
120	2.076	2.196	2.358	2.468	2.617	2.860	3.373			
∞	2.054	2.170	2.326	2.432	2.576	2.807	3.291			