UNIVERSITY OF TORONTO

Faculty of Arts and Science

JUNE EXAMINATION 2004

STA 302 H1F / STA 1001 H1F

Duration - 3 hours

Aids Allowed: Calculator

NAME: ______ SOLUTIONS______

STUDENT NUMBER: _____

- There are 17 pages including this page.
- The last page is a table of formulae that may be useful.
- Tables of the t distribution can be found on page 15 and tables of the F distribution can be found on page 16.
- Total marks: 75

1abcde	1fg	2ab	2cd	3	4

5a	5b	5cd	5ef	6ab	6cde

1. Olympic gold medal performances in track and field improve over time. A regression was run with dependent variable longjump, the winning distance in the long jump (in inches), and independent variable year, the year the Olympics was held after 1900 (counting 1900 as year 0). Data from the Olympics held from 1900 through 1984 were used (some Olympics were missed during the World Wars). Here are the data:

year	0	4	8	12	20	24	28	32	36	48
longjump	282.9	289.0	294.5	299.3	281.5	293.1	304.8	300.8	317.3	308.0
year	52	56	60	64	68	72	76	80	84	

Some output from SAS is given below.

The REG Procedure Descriptive Statistics

			Uncorrected		Standard		
Variable	Sum	Mean	SS	Variance	Deviation		
Intercept	19.00000	1.00000	19.00000	0	0		
year	824.00000	43.36842	49184	747.13450	27.33376		
longjump	5890.81250	310.04276	1833077	370.73440	19.25446		

The REG Procedure Model: MODEL1 Dependent Variable: longjump

			Analvsis o	f Vari	ance				
			Sum	of		Mean			
Source		DF	Squa	res	Sc	uare	FV	alue	Pr > F
Model		1	5054.88	650	5054.8	8650	5	3.10	<.0001
Error		17	1618.33	267	95.1	9604			
Corrected Tot	al	18	6673.21	916					
	Root MSE		9.75	685	R-Squar	e	0.757	5	
	Dependent	Mean	310.04	276	Adj R-S	g	0.743	2	
	Coeff Var		3.14	694	U	-			
			Parameter	Estima	ates				
		Pa	rameter	Sta	ndard				
Variable	DF	E	stimate		Error	t Va	lue	Pr > t	5
Intercep	ot 1	28	3.45427	4.	28064	66	.22	<.000	01
vear	1		0.61308	0.	08413	7	.29	<.000	01

Questions based on this output are on the next 2 pages.

year

(a) (2 marks) Estimate the mean change in the winning long jump distance from one Olympics to the next, assuming that the Olympics are held every 4 years.

4(0.61308) = 2.453

(b) (1 mark) Estimate the variance in winning long jump distance that is not explained by the year the Olympics were held.

95.196

(c) (1 mark) Estimate the percent of total variability in winning long jump distance that is explained by the year the Olympics were held.

75.75%

(d) (1 mark) Estimate the correlation between winning long jump distance and the year the Olympics were held.

$$\sqrt{.7575} = .8703$$

 (e) (2 marks) What is the observed value of the test statistic for the test H₀: β₁ = 0 versus H_a: β₁ > 0? What is the *p*-value for this test? Test statistic: 7.29 *p*-value: < 0.00005 (f) (3 marks) Construct simultaneous 99% confidence intervals for the slope and intercept of the regression line. Using the Bonferroni method, $t_{17,0.005/2} = 3.222$ CI for slope: .61308 $\pm 3.222(.08413) = (.342, .884)$ CI for intercept: 283.454 $\pm 3.222(4.2806) = (269.66, 297.246)$

(g) (5 marks) It has been suggested that the Mexico City Olympics in 1968 saw unusually good track and field performances, possibly because of the high altitude. Construct an appropriate 95% interval for the predicted winning long jump distance in 1968. Do the data support these suggestions? Explain. $\hat{Y}_{68} = 283.454 + .613(68) = 325.14$ $t_{17,.025} = 2.11$

Prediction interval: $325.14\pm2.11(9.757)\sqrt{1+\frac{1}{19}+\frac{(68-43.368)^2}{49184-19(43.368)^2}} = (303.57, 346.71)$ The value in 1968 was 350.5 which is not in the prediction interval. So Mexico City was unusual.

- 2. (8 marks (2 marks each)) Each of the following plots is a scatterplot of a dependent variable versus an independent variable. We wish to study further the relationship between the two variables. Indicate an appropriate linear regression model based on the plot.

Point in bottom right corner is influential. Fit the model $Y_i = \beta_0 + \beta_i X_i + \epsilon_i$ on X in the range from 7.5 to 14.

(b)

(a)



Fit model $Y_i = \beta_0 + \beta_1 X_i + \beta_2 d_i + \beta_3 X_i * d_i + \epsilon_i$ where d_i is 1 if the *i*th point is in one of the groups of data and 0 otherwise.

(This question continues on the next page.)



This shows curvature plus increasing variance so use $Y' = \log(Y)$ or \sqrt{Y} and fit $Y'_i = \beta_0 + \beta_1 X_i + \epsilon_i$

(d)



Fit model $Y_i = \beta_0 + \beta_1 X'_i + \beta_2 (X'_i)^2 + \epsilon_i$ where X'_i is centred X_i

3. (3 marks) Show, in the case of simple linear regression, that the fitted line passes through the point $(\overline{X}, \overline{Y})$. *Fitted line:* $\hat{Y} = b_0 + b_1 X$ where $b_0 = \overline{Y} - b_1 \overline{X}$ *So at* $\overline{X}, \ \hat{Y} = \overline{Y} - b_1 \overline{X} + b_1 \overline{X} = \overline{Y}$

4. (a) (5 marks) A multiple linear regression model is be constructed to examine the relationship between a response variable Y and 3 predictor variables X_1 , X_2 , and X_3 . Suppose measurements of these 4 variables have been taken on n items. State the multiple linear regression model in matrix terms, defining all of your matrices, including the standard assumptions.

Model: $Y = X\beta + \epsilon$ where

$$Y = \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} \quad X = \begin{pmatrix} 1 & X_{11} & X_{12} & X_{13} \\ 1 & X_{21} & X_{22} & X_{23} \\ \vdots & \vdots & \vdots & \vdots \\ 1 & X_{n1} & X_{n2} & X_{n3} \end{pmatrix} \quad \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} \quad \epsilon = \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

Assumptions: $\boldsymbol{\epsilon} \sim N(\mathbf{0}, \sigma^2 \mathbf{I})$

(b) (3 marks) Derive the expression for the covariance matrix of the least squares estimators of the model coefficients of your model in part (a). (I.e., Show $\text{Cov}(\mathbf{b}) = \sigma^2 (\mathbf{X}' \mathbf{X})^{-1}$.)

$$Cov(\mathbf{b}) = Cov((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y})$$

= $(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'Cov(\mathbf{Y})\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}$
= $(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'Cov(\epsilon)\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}$
= $(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\sigma^{2}\mathbf{I}\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}$
= $\sigma^{2}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}$
= $\sigma^{2}(\mathbf{X}'\mathbf{X})^{-1}$

- 5. The data considered in the analysis below are observations on the acceleration (accel) of different vehicles along with their weight-to-horsepower ratio (whp), the speed at which they were travelling (speed), and the grade of the road (grade) which takes values 0, 2, and 6 (a value of 0 indicates the road was horizontal). There are 50 observations in the data set.
 - (a) (5 marks) The first model tried for these data was

 $accel = \beta_0 + \beta_1 whp + \beta_2 speed + \beta_3 grade + \epsilon$

Some output from SAS is given below.

1

grade

The REG Procedure Model: MODEL1 Dependent Variable: accel

			Analysis o	f Varia	ince					
			Sum	of		Mea	an			
Source		DF	Squa	res	S	quai	re	F Val	lue	Pr > F
Model		3	164.994	430	54.	998:	10	25	.45	<.0001
Error		46	99.41	390	2.	161:	17			
Corrected Tot	tal	49	264.40	820						
	Root MSE		1.47	009	R-Squa	re	0	.6240		
	Dependent	Mean	2.60	600	Adj R-	Sq	0	.5995		
	Coeff Var		56.41	183	-	-				
			Parameter 1	Estimat	es					
		Pai	rameter	Star	dard					
Variable	e DF	Es	stimate	E	rror	t	Valu	e I	Pr >	t
Interce	ot 1	7	7.19950	0.6	0087		11.9	8	<.00	01
whp	1	-(0.01838	0.0	0269		-6.8	3	<.00	01
speed	1	-(0.09347	0.0	1307		-7.1	5	<.00	01

What does this output tell you about the ability of the vehicles to accelerate under various conditions? Your answer should explain the affects of each of whp, speed, and grade on acceleration.

-0.15548

0.09040

-1.72

0.0922

For vehicles going the same speed on a road with the same grade, as whp goes up one unit, acceleration goes down 0.0184 units on average. There is strong evidence that the effect on acceleration over and above the other variables is non-zero.

For vehicles with the same whp on roads with the same grade, as speed goes up one unit, acceleration goes down 0.0935 units on average. There is strong evidence that the effect on acceleration over and above the other variables is non-zero.

For vehicles going the same speed on a road with the same whp, as the grade of the road goes up one unit, acceleration goes down 0.155 units on average. There is only weak evidence that this effect is different from zero. (b) (3 marks) The next model fit was

 $\texttt{accel} = \beta_0 + \beta_1 \texttt{whp} + \beta_2 \texttt{speed} + \beta_3 \texttt{grade0} + \beta_4 \texttt{grade2} + \epsilon$

where grade0 = 1 if grade is 0 and is 0 otherwise, and grade2 = 1 if grade is 2 and is 0 otherwise. Some output from SAS for this model follows.

			The REG Proc	edure				
		Depe	endent Variabl	e: accel				
		А	nalysis of Va	riance				
			Sum of		Mean			
		DF	Squares	Sq	uare	F	Value	Pr > F
		4	165.00484	41.2	5121		18.67	<.0001
		45	99.40336	2.2	0896			
Tota	al	49	264.40820					
F	Root I	MSE	1.48626	R-Squar	е	0.62	41	
Dependent Mean		2.60600	Adj R-S	q	0.59	06		
(Coeff	Var	57.03217	-	_			
			Parameter Est	imates				
		Parameter	Standard	l				
Ι)F	Estimate	Error	• t Value	Pr	> t	I	Type I SS
t	1	6.25693	0.61730	10.14		<.000	1	339.56180
	1	-0.01839	0.00272	-6.76		<.000	1	52.93850
	1	-0.09345	0.01322	-7.07		<.000	1	105.66329
	1	0.93165	0.54864	1.70	(0.096	4	3.48062
	1	0.65181	0.56668	1.15	(0.256	1	2.92243
	Tota I (I	Total Root I Depend Coeff DF 1 1 1 1 1	Depe A DF 4 45 Total 49 Root MSE Dependent Mean Coeff Var Parameter DF Estimate 1 6.25693 1 -0.01839 1 -0.09345 1 0.93165 1 0.65181	The REG Prod Dependent Variabl Analysis of Va Sum of DF Squares 4 165.00484 45 99.40336 Total 49 264.40820 Root MSE 1.48626 Dependent Mean 2.60600 Coeff Var 57.03217 Parameter Est Parameter Standard DF Estimate Error 5 1 6.25693 0.61730 1 -0.01839 0.00272 1 -0.09345 0.01322 1 0.93165 0.54864 1 0.65181 0.56668	The REG Procedure Dependent Variable: accel Analysis of Variance Sum of DF Squares Sq 4 165.00484 41.2 45 99.40336 2.2 Total 49 264.40820 Root MSE 1.48626 R-Squares Dependent Mean 2.60600 Adj R-S Coeff Var 57.03217 Parameter Estimates Parameter Standard DF Estimate Error t Value 1 6.25693 0.61730 10.14 1 -0.01839 0.00272 -6.76 1 -0.09345 0.01322 -7.07 1 0.93165 0.54864 1.70 1 0.65181 0.56668 1.15	The REG Procedure Dependent Variable: accel Analysis of Variance Sum of Mean DF Squares Square 4 165.00484 41.25121 45 99.40336 2.20896 Total 49 264.40820 Root MSE 1.48626 R-Square Dependent Mean 2.60600 Adj R-Sq Coeff Var 57.03217 Parameter Estimates Parameter Standard DF Estimate Error t Value Pr 1 6.25693 0.61730 10.14 1 -0.01839 0.00272 -6.76 1 -0.09345 0.01322 -7.07 1 0.93165 0.54864 1.70 1	$\begin{array}{c c} & The \ REG \ Procedure \\ Dependent \ Variable: \ accel \\ \\ & Analysis \ of \ Variance \\ & Sum \ of \\ & Mean \\ DF \\ & Squares \\ & Square \\ & f \\ & 4 \\ & 165.00484 \\ & 41.25121 \\ & 45 \\ & 99.40336 \\ & 2.20896 \\ \hline \\ Total \\ & 49 \\ & 264.40820 \\ \hline \\ Total \\ & 49 \\ & 264.40820 \\ \hline \\ Root \ MSE \\ & 1.48626 \\ & R-Square \\ & 0.62 \\ & Dependent \ Mean \\ & 2.60600 \\ & Adj \ R-Sq \\ & 0.59 \\ & Coeff \ Var \\ & 57.03217 \\ \hline \\ & Parameter \ Estimates \\ \hline \\ Parameter \\ & Standard \\ & DF \\ & Estimate \\ & Error \ t \ Value \ Pr > \ t \\ & 1 \\ & 6.25693 \\ & 0.61730 \\ & 10.14 \\ & <.000 \\ & 1 \\ & -0.01839 \\ & 0.00272 \\ & -6.76 \\ & <.000 \\ & 1 \\ & 0.93165 \\ & 0.54864 \\ & 1.70 \\ & 0.096 \\ & 1 \\ & 0.65181 \\ \hline \end{array}$	$\begin{array}{c c} The REG Procedure \\ Dependent Variable: accel \\ \hline \\ Analysis of Variance \\ Sum of Mean \\ DF Squares Square F Value \\ 4 165.00484 41.25121 18.67 \\ 45 99.40336 2.20896 \\ \hline \\ Total 49 264.40820 \\ \hline \\ Root MSE 1.48626 R-Square 0.6241 \\ Dependent Mean 2.60600 Adj R-Sq 0.5906 \\ \hline \\ Coeff Var 57.03217 \\ \hline \\ \hline \\ Parameter Estimates \\ \hline \\ Parameter Standard \\ DF Estimate Error t Value Pr > t \\ 1 6.25693 0.61730 10.14 <.0001 \\ 1 -0.01839 0.00272 -6.76 <.0001 \\ 1 -0.09345 0.01322 -7.07 <.0001 \\ 1 0.93165 0.54864 1.70 0.0964 \\ 1 0.65181 0.56668 1.15 0.2561 \\ \hline \end{array}$

What is the purpose of using the two variables grade0 and grade2? Can then model the effect of the three grades without making any assumptions about the functional form of the relationship of grade with acceleration. Need 2 indicator variables since there are 3 levels of grade (so a grade of 6 is indicated when both are 0).

Continued

(c) (6 marks) Is the evidence for a grade effect stronger or weaker from the model in part (b) than the model in part (a)? Support your answer with appropriate statistical tests.
First model: p-value for test of no grade effect is 0.0922.
Second model:
Test H₀: β₃ = β₄ = 0 versus H_a: at least one of β₃, β₄ is not zero.
Test statistic: F_{obs} = (3.48+2.92)/2/2.21 = 1.45
Under H₀ this is an observation from an F_{2,45} distribution. Approximating with F_{2,30} gives 0.1
So the evidence of a grade effect is stronger for the first model.

(d) (2 marks) Note that R^2 is almost identical for the models fit in parts (a) and (b). Choose another statistic that is useful for choosing between the two models and indicate which model is preferred.

Adjusted \mathbb{R}^2 : First model is .5995, second model is .5906 so first model is preferred.

OR

MSE: First model is 2.16117, second model is 2.20896 so first model is preferred.

(e) (3 marks) Another model that could be fit is

 $\texttt{accel} = \beta_0 + \beta_1 \texttt{whp} + \beta_2 \texttt{speed} + \beta_3 \texttt{grade} + \beta_4 \texttt{whp_speed} + \epsilon$

where whp_speed = whp*speed. What additional information could be obtained from this model and how would you assess whether or not it is statistically important?

Can see whether the way whp affects acceleration varies with speed.

Test H_0 : $\beta_4 = 0$ versus H_a : $\beta_4 \neq 0$ given other variables in model using usual *t*-test.

(f) (5 marks) A plot of the residuals versus the predicted values and a normal quantile plot of the residuals for the regression in part (a) are on the next page. What additional information do these plots give?

First plot shows increasing variance so need transformation of Y (or weighted least squares).

Normal quantile plot shows tails that are a little lighter than normal so tests and confidence intervals based on normal theory are only approximately correct.



- 6. (a) (6 marks) A simple linear regression is carried out and one point is determined to be an outlier but not an influential point. It is removed from the data and the regression line is fit to the remaining data. Which of the following quantities will differ by a substantial amount between the two regressions? If there is a substantial difference, indicate in which regression it is larger. Indicate if you don't have enough information to say.
 - i. the slope won't differ by a substantial amount since not influential
 - ii. the mean square error will decrease

iii. R^2 will increase

(b) (3 marks) Explain the difference between an outlier and an influential point. An outlier has a large (in absolute value) residual. The point is far from the fitted line. An influential point is a point such that if it is removed and the line re-fit, the line will be different by a substantial amount. A point can be one or both of these.

- (c) (2 marks) If you could choose the values of X at which to collect data before performing a simple linear regression analysis, would you prefer that ∑ⁿ_{i=1}(X_i − X)² be large or small? Explain. Large because gives smaller estimates of the variance of the regression parameter estimates.
- (d) (2 marks) An investigator wishes to use multiple regression to predict a variable, Y, from two other variables, X₁ and X₂. She is also interested in the quantity that is the sum of X₁ and X₂ and includes a third predictor variable in her model, X₃ = X₁ + X₂. What problems might she encounter?
 X'X will be singular so can't calculate regression coefficients
- (e) (4 marks) In a study of infant mortality, a regression model was constructed using birth weight (which is a good indicator of the baby's likelihood of survival) as a dependent variable and several independent variables, including the age of the mother, whether the mother smoked or took drugs during pregnancy, the amount of medical attention she had, her income, etc. The R^2 was 11%, but the *t*-test provided by SAS for the coefficient of each predictor variable had *p*-value less than 0.01. An obstetrician has asked you to explain the significance of the study as it relates to her practice. What would you say to her?

Since each variable has p-value < 0.01, there is strong evidence that each helps explain birth weight over and above the other variables.

For 2 mothers with everything else in the model the same, the coefficient of a variable gives an indication how, on average, the birth weight will change when that variable changes (if positive, good, if negative, bad).

But, since R^2 is low, there is still a lot of variability in the data that is not explained by the model and the birth weight for any given patient may not be well predicted.

Because this is not an experiment we can not say that any of these variables cause a change in birth weight, only that there is an indication of a relationship.