STA 302 / 1001 H - Summer 2004
Test 2
June 16, 2004

LAST NAME: $\qquad$ FIRST NAME: $\qquad$ STUDENT NUMBER: $\qquad$

ENROLLED IN: (circle one) STA 302 STA 1001

INSTRUCTIONS:

- Time: 60 minutes
- Aids allowed: calculator.
- A table of values from the $t$ distribution is on the second to last page (page 8 ).
- A table of formulae is on the last page (page 9).
- Total points: 30

| 1 | 2 | 3 ab | 3 cdef |
| :---: | :---: | :---: | :---: |
|  |  |  |  |

1. Suppose the following four pairs of observations have been made | $Y_{i}$ | 1 | 2 | 1 | 3 |
| :--- | :--- | :--- | :--- | :--- |

| $X_{i}$ | 0 | 1 | 2 | 3 |
| :--- | :--- | :--- | :--- | :--- |

and a simple linear regression is to be fit to the data.

$$
\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-1}=\left(\begin{array}{cc}
7 / 10 & -3 / 10 \\
-3 / 10 & 1 / 5
\end{array}\right) \quad \text { and } \quad \mathbf{e}^{\prime} \mathbf{e}=1.5
$$

(a) (1 point) State the $\mathbf{X}$ matrix.
(b) (3 marks) Find the least squares estimates of the slope and intercept of the regression line.
(c) (2 marks) Estimate the covariance between the estimators of the slope and intercept.
2. For the multiple linear regression model $\mathbf{Y}=\mathbf{X} \beta+\epsilon$, the least squares estimates are $\mathbf{b}=\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-1} \mathbf{X}^{\prime} \mathbf{Y}$ and the residuals are $\mathbf{e}=\mathbf{Y}-\mathbf{X b}$. Assume the GaussMarkov conditions hold.
(a) (3 marks) Show that $\mathbf{b}$ is an unbiased estimator of $\beta$.
(b) (3 marks) You may recall that in simple linear regression, $\sum_{i=1}^{n} X_{i} e_{i}=0$. For multiple linear regression, show that $\mathbf{X}^{\prime} \mathbf{e}=\mathbf{0}$.
3. This question looks at the results of a study of gas chromatography, a technique which is used to detect very small amounts of a substance. Five measurements were taken for each of four specimens containing different amounts of the substance. The amount of the substance in each specimen was determined before the experiment. The response variable is the output reading from the gas chromatograph. Although the observations were made by the same machine over time, assume that they are independent. Some output from SAS is given below.


Questions based on this output start on the next page.
(a) (4 points) Predict the amount of the substance which will give a response of 500 units and construct an appropriate $95 \%$ interval for this prediction.
(b) (2 points) What residual plots would you like to see to check whether it is reasonable to "treat the observations as independent"?

Given below are a plot of the residuals versus the predicted values and a normal probability plot of the residuals for the gas chromatography data.



The questions that follow on the next page relate to these two residual plots.
(c) (2 marks) Describe (with a sketch or in words) what a plot of the residuals versus amount (the independent variable) would look like.
(d) (5 marks) Describe any problems you see with the given residual plots. Indicate what assumptions about the model are being violated.
(e) (2 marks) How do your comments about the residual plots made in the previous part affect the interpretation of your answer to part (a) on page 5?
(f) (3 marks) What would be appropriate action to remedy any problems identified in the residual plots? Justify your answer.

## Some formulae:

$$
\begin{array}{cc}
b_{1}=\frac{\sum\left(X_{i}-\bar{X}\right)\left(Y_{i}-\bar{Y}\right)}{\sum\left(X_{i}-\bar{X}\right)^{2}} & b_{0}=\bar{Y}-b_{1} \bar{X} \\
\operatorname{Var}\left(b_{1}\right)=\frac{\sigma^{2}}{\sum\left(X_{i}-\bar{X}\right)^{2}} & \operatorname{Var}\left(b_{0}\right)=\sigma^{2}\left(\frac{1}{n}+\frac{\bar{X}^{2}}{\sum\left(X_{i}-\bar{X}\right)^{2}}\right)
\end{array}
$$

$$
\operatorname{Cov}\left(b_{0}, b_{1}\right)=-\frac{\sigma^{2} \bar{X}}{\sum\left(X_{i}-\bar{X}\right)^{2}}
$$

$$
\mathrm{SSTO}=\sum\left(Y_{i}-\bar{Y}\right)^{2}
$$

$$
\mathrm{SSE}=\sum\left(Y_{i}-\hat{Y}_{i}\right)^{2}
$$

$$
\mathrm{SSR}=b_{1}^{2} \sum\left(X_{i}-\bar{X}\right)^{2}=\sum\left(\hat{Y}_{i}-\bar{Y}\right)^{2}
$$

$$
\sigma^{2}\left\{\hat{Y}_{h}\right\}=\operatorname{Var}\left(\hat{Y}_{h}\right)
$$

$$
=\sigma^{2}\left(\frac{1}{n}+\frac{\left(X_{h}-\bar{X}\right)^{2}}{\sum\left(X_{i}-\bar{X}\right)^{2}}\right)
$$

$$
\begin{aligned}
\sigma^{2}\{\text { pred }\} & =\operatorname{Var}\left(Y_{h}-\hat{Y}_{h}\right) \\
& =\sigma^{2}\left(1+\frac{1}{n}+\frac{\left(X_{h}-\bar{X}\right)^{2}}{\sum\left(X_{i}-\bar{X}\right)^{2}}\right)
\end{aligned}
$$

$$
\hat{X}_{h} \pm \frac{t_{n-2,1-\alpha / 2}}{\left|b_{1}\right|} * \text { appropriate s.e. }
$$

(valid approximation if $\frac{t^{2} s^{2}}{b_{1}^{2} \sum\left(X_{i}-\bar{X}\right)^{2}}$ is small)

Working-Hotelling coefficient:

$$
W=\sqrt{2 F_{2, n-2 ; 1-\alpha}}
$$

$$
\begin{array}{rlrl}
\operatorname{Cov}(\mathbf{X}) & =E\left[(\mathbf{X}-E \mathbf{X})(\mathbf{X}-E \mathbf{X})^{\prime}\right] \\
& =\mathrm{E}\left(\mathbf{X} \mathbf{X}^{\prime}\right)-(\mathrm{EX})(\mathrm{EX})^{\prime} & \operatorname{Cov}(\mathbf{A X})=\mathbf{A} \operatorname{Cov}(\mathbf{X}) \mathbf{A}^{\prime} \\
\mathbf{b}=\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-1} \mathbf{X}^{\prime} \mathbf{Y} & \operatorname{Cov}(\mathbf{b})=\sigma^{2}\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-1} \\
\hat{\mathbf{Y}}=\mathbf{X} \mathbf{b}=\mathbf{H Y} & \mathbf{e}=\mathbf{Y}-\hat{\mathbf{Y}}=(\mathbf{I}-\mathbf{H}) \mathbf{Y} \\
\mathbf{H}=\mathbf{X}\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-1} \mathbf{X}^{\prime} & \mathrm{SSR}=\mathbf{Y}^{\prime}\left(\mathbf{H}-\frac{1}{n} \mathbf{J}\right) \mathbf{Y} \\
\mathrm{SSE}=\mathbf{Y}^{\prime}(\mathbf{I}-\mathbf{H}) \mathbf{Y} & \mathrm{SSTO}=\mathbf{Y}^{\prime}\left(\mathbf{I}-\frac{1}{n} \mathbf{J}\right) \mathbf{Y}
\end{array}
$$

